

1 **Title: Risk and Resilience Predictors of Recovery after Spinal Fusion**  
2 **Surgery in Adolescents**

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## ABSTRACT

1  
2 **Objective:** This prospective study examined risk and resilience predictors of pain and functional  
3 recovery in the first six months after major surgery in adolescents. **Methods:** Adolescents with  
4 Adolescent Idiopathic Scoliosis undergoing spinal fusion surgery (n = 100, aged 12 to 18 years, 77%  
5 girls) completed assessments prior to surgery, and at three weeks, six weeks, and six months after  
6 surgery. Recovery trajectories in pain, health-related quality of life, and objectively registered physical  
7 activity were identified. Pre-surgical pain catastrophizing and pain intensity (risk), and psychological  
8 flexibility and postsurgical pain acceptance (resilience) were examined as predictors of recovery.  
9 **Results:** Latent growth class analyses revealed four distinct pain recovery trajectories (i.e., *Severe-*  
10 *Moderate* (11 %, n = 9), *Mild-No* (58%, n = 49), *Moderate-Mild* (24%, n = 20), and *Moderate-Severe*  
11 (7%, n = 6) pain trajectory), two Health-Related Quality of Life (HRQOL) recovery trajectories, two  
12 trajectories characterizing recovery in average daily physical activity at moderate-to-vigorous intensity  
13 (MVPA), and three trajectories characterizing recovery in total physical activity volume characterized  
14 by the average daily number of steps. Subsequent MANOVA analyses revealed that pre-surgical pain  
15 intensity (partial  $\eta^2 = .21$ ,  $p < .001$ ) and pain catastrophizing (partial  $\eta^2 = .13$ ,  $p < .01$ ) were both predictive  
16 of poorer recovery in HRQOL, and pain catastrophizing additionally predicted poorer pain recovery  
17 (partial  $\eta^2 = .15$ ,  $p < .05$ ). Psychological flexibility (partial  $\eta^2 = .25$ ,  $p < .001$ ) and postsurgical pain  
18 acceptance (partial  $\eta^2 = .07$ ,  $p < .05$ ) were predictive of more favorable recovery trajectories in HRQOL,  
19 and psychological flexibility additionally predicted more favorable recovery trajectories in postsurgical  
20 pain (partial  $\eta^2 = .15$ ,  $p < .05$ ). Daily MVPA trajectories were not significantly predicted by any of the  
21 hypothesized factors, while pre-surgical pain catastrophizing levels were predictive of a delayed  
22 recovery trajectory in daily amount of steps (partial  $\eta^2 = .17$ ,  $p < .01$ ). **Conclusions:** Pre-surgical  
23 screening could include assessment of pain intensity, pain catastrophizing, psychological flexibility, and  
24 pain acceptance to identify adolescents who are at-risk for poorer recovery. These are potentially  
25 modifiable factors that can be targeted in pre-surgical interventions to prevent poor and foster adaptive  
26 outcomes after major surgery in adolescents.  
27 **Keywords:** risk, resilience, postsurgical, recovery, pain, adolescent idiopathic scoliosis

## INTRODUCTION

1  
2  
3 Spinal fusion surgery is one of the most common major surgical procedures performed in youth <sup>1</sup>  
4 and has been associated with moderate to severe pain levels and significant impairments in daily  
5 functioning <sup>4,14,25</sup>. Although some levels of acute postsurgical pain are normal when recovering from  
6 such major surgery, approximately 20% of children and adolescents continue to show moderate to severe  
7 pain levels up to one year after the surgery <sup>2</sup>. Chronic postsurgical pain (i.e., “pain lasting for three  
8 months or longer after surgery that is not otherwise associated with pre-existing problems or postsurgical  
9 complications” <sup>3</sup> has attracted increasing attention within the pediatric pain literature. Interestingly,  
10 research has shown that the course of postsurgical pain recovery can differ considerably between  
11 adolescents <sup>2</sup>. For instance, one study showed up to five distinct trajectories in pain recovery in a sample  
12 of adolescents who underwent spinal fusion surgery <sup>4</sup>. Undergoing major surgery may likewise cause  
13 significant impairments in other domains of functioning beyond pain. In the first weeks following  
14 surgery most children and adolescents experience decreases in their physical and psychosocial quality  
15 of life <sup>5</sup>, with about 10% still reporting impairments in their quality of life and daily functioning at one  
16 year after surgery <sup>6</sup>.

17  
18 Inter-individual differences in postsurgical health outcomes have motivated researchers to  
19 identify biological, psychological, and social *risk factors* that are predictive of poor recovery. Pre-  
20 surgical pain intensity and anxiety have been identified as the most important predictors of the  
21 development of chronic postsurgical pain in youth undergoing major surgery <sup>1,2</sup>. Although pre-surgical  
22 pain catastrophizing has been identified as an important risk factor in many adult studies<sup>9</sup>, current  
23 pediatric studies are limited and inconclusive with some demonstrating it to predict worse pain  
24 immediate after surgery<sup>10</sup> and others reporting no link with pain recovery<sup>5,8,11</sup>.

25 Despite the progress that has been made in identifying risk factors that predict poorer postsurgical  
26 recovery following major surgery in youth, there are still several empirical, theoretical, and  
27 methodological gaps that can be acknowledged in this area. First, the aforementioned work has largely

1 focused on the study of postsurgical *pain* recovery, whereas relatively less attention has been paid to  
2 other domains of functioning<sup>5-8</sup>. One study has examined predictors of poorer recovery trajectories in  
3 *quality of life*<sup>5</sup>, another has documented different trajectories of recovery in *physical activity* levels in  
4 the initial weeks following in- and outpatient surgery in children and adolescents<sup>8</sup> and one recent study  
5 has examined predictors of functional disability at one year after pediatric surgery<sup>9</sup>. Moreover, most  
6 studies have relied on self-report measures to assess postsurgical recovery in adolescents, although this  
7 subjective procedure is known to be prone to several biases (e.g., self-presentation, memory biases)<sup>24</sup>.  
8 Objective monitoring of pre- and postsurgical physical activity levels (e.g., using accelerometers) has  
9 been suggested as a useful tool to indirectly assess functional recovery in terms of re-engagement in  
10 physical activities after surgery (e.g.,<sup>8</sup>).

11 Second, most of this work has focused on examining risk factors for poorer postsurgical  
12 outcomes<sup>1,2,4,6,11,14</sup>. Yet the majority of children and adolescents return to pre-surgical pain and  
13 functional levels within an expected recovery time<sup>2</sup>. Although identifying and targeting risk factors  
14 seems critical, *protective* factors that predict normal recovery following surgery may be equally  
15 important to consider in pre-surgical prevention strategies. In the wider pain literature a risk-resilience  
16 perspective has been proposed which promotes the study of these so-called ‘resilience’ factors that may  
17 predict adaptive functioning in the presence of pain or physical complaints<sup>15,16</sup>. *Resilience* is defined as  
18 “effective functioning despite the exposure to stressful circumstances and/or internal distress”<sup>16,17</sup>. To  
19 date only one study has hinted at such a potential resilience factor for faster pain recovery after surgery,  
20 i.e., greater pain coping efficacy<sup>14</sup>. Another psychosocial resilience factor that may be promising in this  
21 context is psychological flexibility. *Psychological flexibility* can be described as being aware of, and  
22 open to unwanted and uncontrollable inner experiences, while still being able to act in-line with what  
23 one values in life<sup>18,19</sup>. It encompasses flexibility in dealing with challenges in several domains of daily  
24 life (such as those associated with recovery from surgery). Psychological flexibility and one of its  
25 subcomponents, *pain acceptance*, have been related to beneficial functional outcomes in children and  
26 adolescents who are confronted with (persistent) pain<sup>20-23</sup>. No prior studies have examined  
27 psychological flexibility as a predictor of pain or other health outcomes in the context of major pediatric

1 surgery.

2

3         The current study employed a prospective design (with four time points from before surgery up  
4 to six months following surgery) to assess postsurgical recovery across several domains, from pain, and  
5 quality of life, to objectively measured daily physical activity and its predicting factors. In line with  
6 previous work, a first objective was to explore distinct trajectory patterns of postsurgical recovery in a  
7 sample of patients with adolescent idiopathic scoliosis (AIS) who underwent spinal fusion surgery. A  
8 second objective was to examine factors that were predictive of individual differences in postsurgical  
9 recovery. In addition to previous work that identified *risk factors* predictive of worse outcomes, the  
10 present study included the investigation of *resilience factors* that predict more favorable outcomes after  
11 surgery. Pre-surgical pain intensity and pain catastrophizing were examined as potential risk factors. It  
12 was hypothesized that these factors would predict poorer postsurgical pain, health-related quality of life,  
13 and physical activity recovery trajectories. Pre-surgical psychological flexibility and acceptance of  
14 postsurgical pain were examined as potential resilience factors. It was hypothesized that these factors  
15 would predict more favorable recovery in postsurgical pain, health-related quality of life, and physical  
16 activity.

17

## METHODS

18

### 19 **Participants**

20         Adolescents diagnosed with adolescent idiopathic scoliosis (AIS) who were scheduled for spinal  
21 fusion surgery were screened for study eligibility. They were recruited from four orthopedic units in  
22 hospitals in Belgium (i.e., University Hospital Ghent, University Hospital Antwerp, General Hospital  
23 Saint-Jan Bruges, and University Hospital Leuven) between December 2016 and December 2018.  
24 Eligibility criteria for inclusion were (1) being aged between 12 and 18 years, (2) having an AIS  
25 diagnosis, (3) scheduled for posterior spinal fusion surgery, (4) being able to speak and read Dutch, and  
26 (5) having parental informed consent for their participation. Adolescents were considered as non-eligible  
27 if they had (1) prior spinal fusion surgery or (2) a severe comorbid neurological, developmental or

1 another (mental) health condition. The flowchart in Figure 1 illustrates the recruitment procedure. The  
2 majority of refusals were related to concerns about expected time investment and additional mental load  
3 of participating beyond undergoing the surgery. This multi-centric study was approved by the Central  
4 Ethical Committee of Ghent University with permission of all local ethical committees to collect data  
5 at each site, and was carried out according to the guidelines for Good Clinical Practice (ICH/GCP) and  
6 the declaration of Helsinki for the protection of people who participate in clinical studies. Parental  
7 informed consent and adolescent informed assent were obtained online at the start of the study.

8

## 9 **Procedure**

10 This study is part of a larger longitudinal project , the Postoperative Recovery after Spinal Fusion  
11 Surgery (PR-SF) study (the protocol can be found at: <https://biblio.ugent.be/publication/8578153>). This  
12 study reports on adolescent questionnaire and physical activity data that were collected before (T0) and  
13 at three weeks (T1), six weeks (T2), and six months (T3) after surgery.

14 Between three to one week(s) prior to surgery (T0) adolescents completed questionnaires  
15 assessing sociodemographic and biomedical characteristics, predictive factors (i.e., psychological  
16 flexibility, pain catastrophizing, pre-surgical pain intensity), and baseline levels of the postsurgical  
17 outcome variables (i.e., health-related quality of life and postsurgical pain intensity). Baseline physical  
18 activity was assessed by means of a waist-worn accelerometer during seven consecutive days. At three  
19 weeks after surgery (T1) adolescents were asked to complete a questionnaire assessing postsurgical pain  
20 acceptance. Postsurgical outcome variables were assessed using the same questionnaires as at T1, T2,  
21 and T3. Postsurgical physical activity was only monitored at T1 and T2.

22

23 All data were collected at home; participation in this study did not require any additional hospital  
24 visits. Adolescents were reminded about an upcoming time point via telephone calls one week before  
25 they were scheduled start. Questionnaires were completed via an online, secured survey tool (i.e.,  
26 Limesurvey, version 2.00). At each time point an e-mail containing the link and a personal token to  
27 access the surveys was sent to the adolescents. They were asked to complete the surveys within one  
28 week after receiving this e-mail; weekly reminders were sent up to three weeks after the first e-mail. All

1 data collected after these three weeks were considered as invalid and excluded from further analyses.  
2 Accelerometers were provided to the participants during a house visit by a research assistant before  
3 surgery. Once all physical data collection moments were completed, they were asked to return the device  
4 to the research team either via the courier service or during one of their regular hospital appointments.  
5 Biomedical information was collected from the electronic medical record. All adolescents received  
6 standard pre-, peri-, and postsurgical treatment following the regular protocol in each hospital.

7  
8 Participants were given the opportunity to either start their participation at T0 or drop in at a later  
9 point in time and start their participation at T1. This planned missing data design intentionally collects  
10 incomplete data from participants to reduce costs and burden for the participants and has limited effects  
11 on power or bias. In particular, we have accounted for missing data by collecting supplemental entrants  
12 following the same inclusion criteria (i.e. ‘refreshment’ sample)<sup>26,27</sup>. Adolescents who did not complete  
13 data at one time point were allowed to drop back in at a next time point after surgery. Information about  
14 drop-out and drop-in at each time point is presented in Figure 1. Ninety-two adolescents were enrolled  
15 in the study at T0, eight adolescents additionally decided to drop-in at T1. Main reasons for this delayed  
16 start (i.e., after surgery) were concerns about expected time investment and mental load of participating.  
17 Drop-out at subsequent time points was mainly related to adolescents’ lack of time, loss of interest in  
18 the study, or inability of the research team to contact participants. As a retention strategy adolescents  
19 received reminders at each time point and a movie ticket at T3 if all required assessments were  
20 completed and the accelerometer was returned to the research team.

21

22

-INSERT FIGURE 1 ABOUT HERE-

## 23 **Measures**

24 *Sociodemographic and Biomedical Information.* Sociodemographic data such as age, sex,  
25 educational level, and ethnicity were collected by means of a short survey completed prior to surgery.  
26 Biomedical information (i.e., height, weight, Cobb Angle) and length of hospital stay were collected

1 from the electronic medical record by a medical staff member in each hospital at three or six weeks after  
2 surgery.

3  
4 ***Health-Related Quality of Life (T0 – T3).*** Health-related quality of life (HRQOL) was assessed  
5 by means of the child (8-12 years) and teen (13-18 years) versions of the Pediatric Quality of Life  
6 Inventory (PedsQL™ 4.0) <sup>28</sup>. The 23-item PedsQL™ contains five subscales designed to assess  
7 problems with physical, emotional, social, psychological, and school functioning. Items are rated on a  
8 five-point response scale ranging from 0 (*never*) to 4 (*almost always*). A total HRQOL score is obtained  
9 by reverse-scoring and transforming raw scores into standard scores on a 0-100 scale, and subsequently  
10 taking the mean of all item scores. Higher total scores indicate better HRQOL. Scores below the cut-off  
11 of 69.7 are used to identify adolescents that are ‘at-risk’ for impaired HRQOL. This cut-off score was  
12 based on a study comparing PedsQL™ total scores between samples of healthy children and children  
13 with chronic health conditions <sup>28</sup>. The PedsQL™ showed good psychometric properties in population  
14 and clinical samples <sup>28,29</sup>. In the present study Cronbach’s alpha of the PedsQL™ scale (teen version)  
15 ranged between .89 and .97 across the four time points. All adolescents completed this teen version with  
16 the exception of two adolescents who completed the child version at T0 as they were under the age of  
17 13. Cronbach’s alpha for this child scale was .92.

18  
19 ***Physical Activity (T0 – T2).*** At T0, T1, and T2 adolescents wore an accelerometer (i.e.,  
20 ActiGraph™ (GT3X+); Pensacola, Florida) over a period of seven consecutive days to register daily  
21 physical activity. The Actigraph™ is a small, non-invasive device that is worn around the waist.  
22 Adolescents were asked to wear the device at the right side of the body during day and night and to  
23 remove it only for showering, swimming or activities that involve the risk to potentially damage the  
24 device (e.g., fighting sports). Verbal instructions on how to wear the device were given by a member of  
25 the research team during the house visit prior to the start of the study and additionally summarized on  
26 an information leaflet. Actigraph™ devices have been shown to be valid, objective measures to  
27 characterize physical activity levels in adolescents and children in the general population <sup>30</sup> as well as  
28 in clinical <sup>31–33</sup> and postsurgical settings <sup>8</sup>. Both physical activity at moderate-to-vigorous intensity levels



1 (MVPA) and number of steps (i.e., total physical activity volume) were included as indicators of  
2 physical activity recovery.

3

4 ***Pain Intensity (T0 – T3)***. The child version of the Graded Chronic Pain Scale (GCPS-C)<sup>34,35</sup> was  
5 used to measure both pre-surgical pain intensity (T0) and intensity of postsurgical pain at three weeks,  
6 six weeks, and six months after surgery (T1 – 3). Current, worst, and average pain intensity during the  
7 previous three weeks was rated on an 11-point numerical rating scale (0 = *no pain*, 10 = *worst possible*  
8 *pain*) and used to calculate a characteristic pain intensity score at each time point. Note that the time  
9 frame of the original GCPS-C was adjusted from six months to three weeks for use in this longitudinal  
10 study. Following prior work with numerical pain rating scales<sup>36</sup>, scores  $\geq 1$  and  $< 4$  were classified as  
11 *Mild Pain*, scores  $\geq 4$  and  $< 7$  as *Moderate Pain*, and scores  $\geq 7$  as *Severe Pain*. The GCPS-C has been  
12 used as a valid measure of pain severity in primary care, chronic pain, and general population samples  
13<sup>37–39</sup>. The child version has shown good psychometric properties in a general population sample<sup>35</sup>.  
14 Numerical rating scales have demonstrated good validity to measure acute postoperative pain levels in  
15 children and adolescent and is sensitive to measure change in pain intensity over time<sup>40</sup>. Cronbach's  
16 alpha of the pain intensity subscale in this study ranged between .74 and .89 across the different time  
17 points.

18

19 ***Pain Catastrophizing (T0)***. Adolescents completed the Pain Catastrophizing Scale for Children  
20 (PCS-C)<sup>41</sup> before surgery to measure their level of catastrophizing thoughts about pain. The PCS-C  
21 consists of 13 items that are scored on a scale from 0 (*not at all*) to 4 (*extremely*). The PCS-C yields  
22 three subscale scores: rumination (4 items), magnification (3 items), and helplessness (6 items) and a  
23 total score. Higher total scores (0-52) indicate greater catastrophizing and worrying about pain. Good  
24 psychometric qualities of the PCS-C have been shown with healthy and clinical samples<sup>41</sup>. In the current  
25 study a Cronbach's alpha of .97 was found for the PCS-C.

26

27 ***Psychological Flexibility (T0)***. The 17-item Avoidance and Fusion Questionnaire for Youth  
28 (AFQ-Y)<sup>42</sup> was used to assess adolescents' level of psychological flexibility before surgery. The AFQ-

1 Y was originally designed to assess psychological *inflexibility* through assessment of two core  
2 components: experiential avoidance (i.e., avoidance of negative or unwanted experiences) and cognitive  
3 fusion (i.e., being entangled with the content of one's thoughts or feelings). Items are rated on a five-  
4 point response scale from 0 (*not at all true*) to 4 (*very true*). In line with previous research<sup>43,44</sup>, items  
5 were reverse-scored so that higher total scores (0-68) indicate a greater level of psychological *flexibility*.  
6 The AFQ-Y showed good psychometric properties in healthy and clinical samples<sup>42,45</sup>. Cronbach's  
7 alpha of the AFQ-Y scale was .90 in the current study.

8  
9 ***Postsurgical Pain Acceptance (T1)***. Adolescents completed the Chronic Pain Acceptance  
10 Questionnaire for Adolescents (CPAQ-A)<sup>46</sup> at three weeks after surgery. The reason for assessing this  
11 individual factor at T1 instead of at baseline was that it was expected that not all participants would be  
12 in pain before surgery which made it only possible to measure this process after surgery when all  
13 participants would have (had) experienced some pain as a consequence of the invasive surgery. The  
14 CPAQ-A is a 20-item scale that measures acceptance of pain. Items are scored on a five-point response  
15 scale from 0 (*never true*) to 4 (*always true*). The scale contains two subscales assessing Activity-  
16 engagement (i.e., the extent to which adolescents attempt to participate in daily activities when in pain;  
17 11 items) and Pain Willingness (i.e., the extent to which adolescents rate the goal to control their pain  
18 level as less important than other life goals; 9 items). Higher total scores (0-80) indicate greater levels  
19 of pain acceptance. The CPAQ-A has demonstrated good reliability and validity in samples of  
20 adolescents with (chronic) pain<sup>46,47</sup>. In the current study a Cronbach's alpha of .79 was found for the  
21 total CPAQ-A scale.

22

### 23 **Data Preparation**

24 Descriptive and missing data analyses (i.e., Little's MCAR test) were performed using SPSS  
25 (version 25, IBM Statistics). Normality and reliability checks were performed. ActiLife software  
26 (version 6.13.3) was used to transform and analyze raw physical activity data collected during waking  
27 hours. In short, activity counts were calculated based on the adolescent's movements within a pre-  
28 determined time period called an "epoch". In this study epochs of 15 seconds were used (as

1 recommended for children and adolescents by <sup>48</sup>). If a period of 20 consecutive minutes of zero counts  
2 was encountered this was considered as non-wear time and these data were excluded from further  
3 analyses. Adolescents were asked to wear the accelerometer for at least 8 hours per day to reliably  
4 estimate their physical activity levels on that day. *Evenson* cut-points were used to interpret counts and  
5 estimate the total minutes spent in activities of moderate-to-vigorous physical activity (MVPA) and  
6 number of steps per day <sup>48</sup>. In this study the daily number of steps and minutes of MVPA were  
7 aggregated over the seven-day period to obtain average daily physical activity scores. Data of  
8 participants with less than 4 out of 7 registration days were considered as invalid and excluded from  
9 further analyses. The World Health Organization (WHO) recommends that all children and adolescents  
10 should participate in at least 60 minutes of MVPA on a daily basis to experience physical and  
11 psychological health benefits <sup>49,50</sup>, data were evaluated relative to this guideline to interpret recovery  
12 patterns.

13

#### 14 **Data-analytical strategy**

15 Pre-surgical pain intensity, pain catastrophizing, and psychological flexibility at T0, and pain  
16 acceptance at T1 were included as predictor variables. Postsurgical pain intensity, HRQOL, and physical  
17 activity (i.e., average daily MVPA and step count) were examined as outcome variables. Pearson  
18 product-moment correlations between predictors and outcomes were evaluated at each time point.  
19 Independent sample t-tests were performed to evaluate differences in key outcome variables and/or  
20 sociodemographic characteristics between those who enrolled before surgery and those who “dropped  
21 in” at T1.

22 Next, to answer the research questions, a data-analytical strategy described by Jung & Wickrama  
23 <sup>51</sup> was followed. First, to identify classes of individuals (cfr. Research Question 1), Latent Class Growth  
24 Analysis (LCGA) was performed in Mplus 7 <sup>52</sup>. LCGA is a person-centered approach used to examine  
25 patterns in outcomes over time (i.e., trajectories) and group individuals into classes based on similar  
26 patterns. Although part of growth mixture modeling, LCGA is distinct given that variances and  
27 covariances within each class are fixed to zero, assuming homogeneity of individual growth trajectories  
28 <sup>51</sup>. In the current study, trajectory classes were examined for each outcome separately from prior to

1 surgery (T0) over three weeks (T1), to six weeks (T2) and ending at six months after surgery (T3; not  
2 for physical activity outcomes). An exploratory *class enumeration process* was followed to decide on  
3 how many classes to retain in the final model. This process started with specifying a baseline one-class  
4 model followed by estimating additional models thereby increasing the number of classes. Various fit  
5 indices were checked to compare model fit. First, lower values on the Akaike Information Criterion  
6 (AIC), the Bayesian Information Criterion (BIC), and the sample-size adjusted Bayesian Information  
7 Criterion (aBIC) indicated a better model fit (e.g., <sup>53,54</sup>). Second, entropy values higher than .80 indicated  
8 an adequate probability of group membership <sup>53</sup>. Third, the Vuong-Lo-Mendell-Rubin likelihood ratio  
9 (VLMR-LRT) and Bootstrap Likelihood Ratio test (BLRT) were used to compare models, with a  
10 significant p-value indicating a superior fit of the current  $k$  class compared to the  $k - 1$  class solution.  
11 The model estimation process was terminated if both likelihood ratio tests were insignificant, if the  
12 model did not converge or could not be identified <sup>53,55</sup>. The final model was chosen by (1) comparing  
13 all models' obtained fit indices, (2) plotting these fit indices to detect an "elbow" (i.e., the point from  
14 which there are only small decreases with respect to the fit index<sup>1</sup>) in the model fit across the class  
15 solutions <sup>53</sup>, (3) examining whether there were sufficient (i.e.,  $\geq 5\%$ ) individuals in each class within the  
16 class solution, and (4) by reflecting on the theoretical meaning of each class solution. In a next step, the  
17 effect of potential predefined confounders (i.e., age, sex, and cobb angle) on the final class solution was  
18 explored. If significant, they were entered as control variables in the final model. Finally, trajectories  
19 (i.e. changes in estimated marginal means across time points) were plotted to evaluate and describe the  
20 classes for each outcome.

21 Next, MANOVAs were performed in SPSS to examine whether pre-surgical pain intensity, pain  
22 catastrophizing, psychological flexibility (T0) and postsurgical pain acceptance (T1) were associated to  
23 class membership (cfr. Research Question 2). Specifically, all predictors were entered as dependent  
24 variables, while class membership (for pain intensity, HRQOL, MVPA, and step count) was entered as  
25 an independent variable. Subsequently, in case of a significant multivariate effect, univariate main

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<sup>1</sup> As it is not uncommon for these fit indices to decrease when more classes are added, an inspection of the size of decrease in the fit is recommended <sup>53</sup>.

1 effects were evaluated to examine hypothesized relations between predictors and outcomes. All tests  
2 were evaluated against a 5% significance level. Additionally, a Bonferroni correction was applied to  
3 account for the potential impact of multiple testing on the results ( $p \leq \alpha/15 = 0.003$ ). Higher levels of  
4 pre-surgical pain catastrophizing and pain intensity were expected to predict membership to the delayed  
5 recovery classes, while psychological flexibility and acceptance of postsurgical pain were expected to  
6 predict membership to the more favorable recovery classes.

7 Although no specific power calculation was performed for this study, a minimal sample size of  
8 100 participants was recommended based on the results of a Monte Carlo simulation study examining  
9 the effects of sample size on model convergence, fit indices and parameter estimates in LGCM<sup>56</sup>.

10

11

## RESULTS

### 12 Sample Characteristics

13 The sample included 100 adolescents ( $M_{age} = 15.19$  years ( $SD = 1.55$ ), *range*: 12 – 18) with AIS  
14 who underwent posterior spinal fusion surgery. All sociodemographic and biomedical characteristics of  
15 the final sample are summarized in Table 1. No significant group differences were found with regard to  
16 the main outcome variables (i.e., pain intensity, HRQOL, MVPA and steps) and sociodemographic  
17 variables (i.e., age, sex, and education) between those who enrolled before surgery and those who  
18 “dropped in” at T1. Prior to surgery, 71% ( $n = 58$ ) of the adolescents reported some pain (scores higher  
19 than 0), 57% ( $n = 41$ ) reported moderate pain intensity levels (scores of 4 or higher). Adolescents most  
20 frequently reported back pain (41%), followed by pain in the neck (5%), joint pain (4%), headaches  
21 (3%), abdominal pain (3%), and non-surgery related injury-related wound pain (1%) as their primary  
22 pain location. Mean reported pain intensity prior to surgery was 3.42 ( $SD = 2.80$ ) and ranged between 0  
23 and 8.

24 The average missingness rate across all included variables and different time points was 21%  
25 (range 11 – 37%). On average 6 days (range 4 – 7 days) of usable physical activity data were available  
26 for further analyses. As Little’s MCAR test indicated that the missing data were missing completely at

1 random,  $\chi^2(449) = 458.896, p = .363$ , we used pairwise deletion in SPSS and Robust Full Information  
2 Maximum Likelihood (FIML) estimation in Mplus to handle the missing data.

3

4

- INSERT TABLE 1 HERE -

## 5 Hypothesis Testing

6 **Pain trajectories.** Following the decision making process described in the data analytical strategy,  
7 a four-class model controlling for Cobb Angle was chosen as the optimal one. This four-class solution  
8 yielded the most optimal fit indices and highest entropy value (Table 2 and Figure 2). Furthermore, the  
9 VLM-LRT and BLRT indicated it was better than a three-class solution and, finally, the four-class  
10 solution had the most meaningful interpretation. Exploration of potential confounders showed that the  
11 Cobb angle was significantly related to the intercepts ( $b = -0.01, SE = 0.004, p < 0.001$ ) and slopes ( $b =$   
12  $-0.01, SE = 0.001, p < 0.01$ ) of each of the four pain classes, whereas there was no effect of age and sex.

13 A *Severe-Moderate*, a *Mild-No*, a *Moderate-Mild*, and a *Moderate-Severe* pain class were derived  
14 based on visual inspection of the graph (Figure 3). This graph shows a small increase ( $\pm 1$  point) in pain  
15 intensity from prior surgery to three weeks after surgery (T1) in all classes except for the *Severe-*  
16 *Moderate* pain class. The majority of the sample was classified into the *Mild-No* pain class (58%,  $n =$   
17 49); Figure 3 shows mild pain levels before surgery, which increased at three weeks and subsequently  
18 decreased at six weeks after surgery staying in the mild pain range in this class. At six months after  
19 surgery this class showed complete recovery from pain. Inspection of the graph (Figure 3) furthermore  
20 shows that adolescents in the *Moderate-Mild* pain class (24%,  $n = 20$ ) reported moderate pain intensity  
21 levels before surgery, which increased at three weeks but continually decreased to mild pain levels at  
22 six weeks and six months after surgery. Figure 3 also demonstrates that the *Severe-Moderate* pain class  
23 (11 %,  $n = 9$ ) evidenced a decrease in pain intensity levels at three weeks which continually decreased  
24 at T2, but remained stable after that and still showed moderate pain intensity levels at T3. Finally, the  
25 *Moderate-Severe* pain class (7%,  $n = 6$ ) consistently reported moderate to severe pain intensity levels  
26 (i.e., scores ranging between 5 and 7) at every time point; Figure 3 shows a small decrease from three  
27 weeks to six weeks but increased levels at six months in this class. Pain intensity levels in this group

1 were observed to be higher at six months than those reported prior the surgery and close to the severe  
2 pain cut-off.

3 Results of the MANOVA showed a significant multivariate relation between the four identified  
4 pain trajectories and baseline scores of pain catastrophizing and pain intensity (risk), and psychological  
5 flexibility and pain acceptance (resilience), Wilks'  $\lambda = .745$ ,  $F(9, 138.874) = 1.989$ ,  $p = 0.05$ , partial  $\eta^2$   
6 = .09. More specifically, univariate main effects indicated that class membership was only significantly  
7 related to baseline pain catastrophizing and psychological flexibility (Table 3). Pairwise comparisons of  
8 estimated marginal means showed that baseline pain catastrophizing was significantly higher in the  
9 *Severe-Moderate* pain class than in the *Moderate-Severe* ( $p = 0.019$ ) and the *Mild-No* pain class ( $p =$   
10  $0.019$ ). Furthermore, baseline psychological flexibility levels were higher in the *Mild – No* pain class as  
11 compared to the *Moderate-Mild* ( $p = 0.020$ ) and the *Moderate-Severe* pain class ( $p = 0.015$ ) (Table 3).  
12 Our hypotheses were thus partly confirmed: pain catastrophizing predicted membership to worse pain  
13 recovery trajectories while psychological flexibility predicted more favorable pain recovery.

14  
15 ***Health-related Quality of Life trajectories.*** Following the same decision strategy a two-class  
16 solution was chosen as the most optimal one (Table 2 and Figure 2). This two-class solution had the  
17 highest entropy, the VLM-LRT indicated it to be better than the one-class solution and it had the most  
18 meaningful interpretation. Age, sex, and Cobb angle were not associated with the intercept and/or slope  
19 of the classes. The final plot of HRQOL trajectories is shown in Figure 3. The *At-Risk* class (32%,  $n =$   
20 32) describes a group of adolescents who were at risk for impaired HRQOL (i.e., with HRQOL below  
21 the at-risk cut-off of 69.70<sup>57</sup>) before surgery. Visual inspection of the graph shows a decrease in HRQOL  
22 at T1 in this group. Although HRQOL continually improved from T1 to T2, and from T2 to T3, this  
23 group still showed “at-risk” HRQOL levels at six months after surgery, basically reaching similar  
24 HRQOL level as before surgery. Figure 3 furthermore shows that the *Resilient* class (68 %,  $n = 65$ )  
25 started at healthy levels of HRQOL before surgery (i.e., above the at-risk cut off), which also decreased  
26 at three weeks (T1) after the surgery. Similar to the other class, the trajectory graph shows continual  
27 improvements in HRQOL levels at six weeks and six months after surgery in this group, returning to  
28 pre-surgical “healthy” HRQOL levels at six months.

1 Results of the MANOVA including all variables show a significant multivariate effect of class  
2 membership, Wilks'  $\lambda = .630$ ,  $F(4,66) = 9.681$ ,  $p < .001$ , partial  $\eta^2 = .37$ . Subsequent univariate analyses  
3 indicated significant main effects of class membership on pre-surgical pain intensity, pain  
4 catastrophizing, psychological flexibility, and postsurgical pain acceptance levels (Table 3). Pairwise  
5 comparison of estimated marginal means showed that pre-surgical pain intensity and pain  
6 catastrophizing were lower, and psychological flexibility and postsurgical pain acceptance were higher  
7 in the *Resilient* class as compared to the *At-Risk* class (Table 3). This supports the hypotheses that pre-  
8 surgical pain intensity and pain catastrophizing are related to delayed postsurgical recovery in HRQOL,  
9 and that psychological flexibility and pain acceptance are associated with a more favorable recovery in  
10 HRQOL.

11

### 12 ***Physical activity trajectories.***

13 *MVPA trajectories.* Based on the decision strategy resulted a two-class solution was chosen as  
14 the most optimal one. Fit indices showed smaller decreases after both the two- and three-class solutions  
15 (Figure 2), BLRT also indicated that these class solutions were a better fit, and entropy values were only  
16 found to be adequate ( $>.80$ ) for the three-class solution (Table 2). Yet, the two-class solution was  
17 retained because of its more favorable theoretical meaning and because the potential third class included  
18 less than 5% of the sample. Age, sex, and Cobb angle were no significant confounders of the intercepts  
19 and/or slopes of the final two trajectory classes.

20 Figure 3 shows that there was only a small difference between the trajectory classes in average  
21 daily minutes MVPA prior to surgery (i.e., 33.71 vs. 29.88 minutes), which is in both classes below the  
22 healthy cut-off of 60 minutes of MVPA on a daily basis for youth (as recommended by the World Health  
23 Organization (2011)). Further visual inspection of the trajectory graph of the *Low MVPA* class (11%,  $n$   
24 = 12) shows a small decrease in average daily minutes of MVPA at three weeks, which seem to improve  
25 in the direction of pre-surgical levels at six weeks after surgery. The trajectory graph of the *Very Low*  
26 *MVPA* class (89%,  $n = 78$ ) showed a bigger decrease to less than 10 average daily minutes of MVPA at  
27 three weeks surgery, which showed only small increases at six weeks after surgery.



1 MANOVA results show no significant multivariate effect of class membership on the  
2 hypothesized risk and resilience predictors, Wilks'  $\lambda = .964$ ,  $F(4,61) = .562$ ,  $p = 0.691$ , partial  $\eta^2 = .04$ .  
3 Further inspection of univariate main effects similarly showed no significant differences between the  
4 classes in terms of pre-surgical pain intensity, pain catastrophizing, psychological flexibility, or pain  
5 acceptance (Table 3). In contrast to what was hypothesized, these predictors were not associated with  
6 membership to the *Low* or *Very Low MVPA* class.

7 *Steps trajectories.* A three-class solution, controlling for baseline age, was retained as the most  
8 optimal and meaningful solution based on visual inspection of decreases in fit indices (Figure 2), and  
9 the BLRT, which showed a better fit for the three- as compared to the two-class solution (Table 2).  
10 Baseline age significantly affected intercept and slopes of two classes and were added to the final three-  
11 class model (Table 2).

12 The graph in Figure 3 was used to interpret and label each of the classes. All three classes showed  
13 a significant decrease in average daily step count at three weeks after surgery (T1) and an increase from  
14 three to six weeks (T2). The classes differ in baseline levels of step count and following recovery  
15 trajectory. Visual inspection of the graph shows a *Declined Steps* class with the majority of participants  
16 (48%,  $n = 37$ ), characterized by the highest baseline level of steps and a steep decrease at T1 and T2  
17 with the number of steps still being significantly lower than baseline levels at T2. Furthermore, two  
18 "resilient" classes are derived, showing a less steep decrease in average daily steps after surgery and  
19 total number of steps that are not significantly different from baseline at T2 (= recovery). The *Resilient*  
20 *Low Steps* class (13%;  $n = 10$ ) starts at the lowest number of steps before surgery, while the *Resilient*  
21 *High Steps* class (39%,  $n = 29$ ) starts at a higher number of steps at baseline.

22 MANOVA results showed a significant multivariate effect of class membership on all included  
23 risk and resilience variables, Wilks'  $\lambda = .768$ ,  $F(8, 120) = 2.119$ ,  $p = 0.039$ , partial  $\eta^2 = .12$ . Subsequent  
24 univariate analyses indicated one significant main effect of class membership on baseline pain  
25 catastrophizing levels but not on all other variables (Table 3). Pairwise comparison of estimated  
26 marginal means showed that pain catastrophizing levels were significantly higher in the *Declined Steps*  
27 class as compare to the *Resilient High Steps* class (Table 3). Results only give partial support for our  
28 hypotheses, i.e. pain catastrophizing levels might increase the risk for delayed recovery in physical

1 activity levels but pain intensity, psychological flexibility and pain acceptance were no significant  
2 predictors.

3

4 -INSERT TABLE 2, TABLE 3, FIGURE 2 & FIGURE 3 ABOUT HERE-

## 5 **DISCUSSION**

6 Using a prospective longitudinal design, the general aim of this study was to examine recovery  
7 trajectories in pain, HRQOL, physical activity levels after spinal fusion surgery in adolescents with AIS  
8 and identify risk (i.e., pre-surgical pain intensity and pain catastrophizing) and resilience (i.e.,  
9 psychological flexibility and postsurgical pain acceptance) predictors of the different recovery  
10 trajectories.

11

### 12 **Recovery trajectories following spinal fusion surgery**

13 Consistent with previous work <sup>2,4,8,59</sup>, it was found that the largest group of adolescents reported  
14 increased pain intensity levels in the first three weeks after surgery which continually decreased to low  
15 to no pain at six months post-surgery. However, three other (smaller) subgroups which showed worse  
16 pain recovery trajectories were also identified. In line with previous work these delayed recovery groups  
17 consistently showed higher pain intensity levels over the course of six months as compared to the first  
18 group <sup>2,8,60</sup>. Our *Moderate-Severe Pain* group can be compared to a similar worse pain recovery group  
19 in a longitudinal study by Sieberg et al. <sup>4</sup> that examined long-term pain trajectories up to five years after  
20 spinal fusion surgery. Different from this previous study<sup>4</sup>, two novel delayed pain recovery trajectories  
21 emerged in our work. Both groups followed a similar pain recovery course as the majority of our sample  
22 but adolescents in these groups already reported higher pain intensity levels prior to surgery.  
23 Adolescents in the *Moderate-Mild* and the *Severe-Moderate Pain* groups respectively reported  
24 moderate/severe pain levels before surgery which continually decreased to mild/moderate pain levels at  
25 six months after surgery.

26 Given that recovery was only examined in the first six months after surgery it could not be

1 analyzed if pain also eventually disappeared at later times. Although spinal fusion surgery is primarily  
2 executed to correct the spinal curvature and not to resolve pain <sup>1</sup>, our findings suggest that at least some  
3 of the pain experienced prior to surgery was resolved by the surgery. However, this idea remains  
4 speculative as it was not possible to identify whether pain experienced after surgery was caused by the  
5 surgery itself or other (pre-surgical) factors based on these findings. In recent years, the importance of  
6 studying and effectively treating chronic postsurgical pain in children and adolescents has been  
7 emphasized <sup>5</sup>. The current study highlights distinct trajectories in pain recovery after spinal fusion  
8 surgery in adolescents and showed that a subgroup reported moderate to severe chronic pain levels at  
9 six months after surgery. These numbers are in line with what is generally found in terms of adolescent  
10 pain recovery across different types of major surgery <sup>5</sup>.

11

12         Whereas previous work mainly focused on the study of postsurgical pain recovery and the  
13 development of chronic pain, we also examined recovery trajectories in HRQOL (i.e., psychosocial and  
14 physical) and physical activity.

15         Two HRQOL groups emerged that both showed a significant decline in HRQOL in the acute  
16 period after surgery. Similar acute decreases in HRQOL were reported in other studies across different  
17 surgical procedures <sup>7,59</sup>. The critical difference between the two groups was that they already showed  
18 significantly different HRQOL levels before surgery. The largest *Resilient* group showed similar  
19 average HRQOL levels before surgery as compared to a normative sample of healthy adolescents <sup>57</sup>.  
20 This group only reported significant impairments in HRQOL at three weeks after surgery, but bounced  
21 back to healthy HRQOL levels at six months after surgery. The smaller *At-Risk* group already showed  
22 impaired HRQOL levels before surgery when compared to the same healthy norm group. This group  
23 continued to report impaired HRQOL levels across the entire postsurgical period. The few studies that  
24 have documented long-term HRQOL outcomes after major spinal surgery suggest that HRQOL should  
25 be completely restored to healthy levels at one and two year follow-up <sup>8,61</sup>. Our results might point to  
26 the potential existence of a subgroup of adolescents with AIS that may be at risk for long-term impaired  
27 HRQOL after spinal fusion surgery, although this should be further examined in a study including long-  
28 term assessments beyond six months.

1 Finally, postsurgical recovery in physical activity was operationalized by average daily physical  
2 activity at moderate-to-vigorous intensity levels (MVPA) and number of steps. With an average of about  
3 30 minutes of MVPA per day before surgery, none of the adolescents in this sample managed to achieve  
4 the recommended guideline of at least 60 minutes<sup>58</sup>. Research suggests that only 9-25% of Belgian  
5 youth actually reaches the recommended guideline<sup>62</sup>. Furthermore, girls and older children generally  
6 show lower compliance to these physical activity recommendations<sup>63</sup>. This might partly explain the  
7 poor levels that were observed in our predominantly female and adolescent sample. Potentially,  
8 prevalence rates are even lower in this sample of adolescents with AIS because of their health condition.  
9 Two MVPA classes emerged that did not differ in the average daily amount of time doing moderate to  
10 vigorous physical activities before surgery. Yet, the two classes clearly showed a distinct trajectory in  
11 physical activity recovery in the first six months after spinal fusion surgery. While the largest *Very Low*  
12 *MVPA* class showed a decrease to less than 5 minutes MVPA per day at three weeks after surgery which  
13 slightly increased to 10 minutes at six weeks after surgery, the other, smaller, *Low MVPA* class showed  
14 smaller decreases at three weeks which were already restored to pre-surgical levels (i.e. 30 minutes) at  
15 six weeks after surgery. Although not many other studies have examined adolescent physical activity in  
16 a postsurgical context, the findings of the present study resemble with what was found in a recent  
17 longitudinal study that also showed significant declines in the average amount of daily MVPA in the  
18 first weeks after inpatient surgery<sup>13</sup>. Other, cross-sectional, studies similarly demonstrated that  
19 adolescents' level of MVPA was still at 50% of the recommended guideline at one to two years after  
20 other types of major surgery<sup>64,65</sup>. Orthopedic surgeons generally advice their patients to avoid bicycle  
21 riding and swimming in the first six months after spinal fusion surgery. Given that these activities are  
22 typically performed at moderate-to-vigorous intensity, this might partly explain why such low levels are  
23 observed in this sample. Therefore we additionally examined the average number of daily steps as an  
24 indicator of post-surgical changes in total physical activity levels (independent from intensity levels).  
25 Here, three distinct classes were identified, indeed showing greater variation in total physical activity  
26 levels than when only looked at MVPA. The largest group showed a steep decrease in the average daily  
27 amount of steps from baseline to three weeks after surgery which was not yet restored to baseline levels  
28 at six weeks after surgery. Two other "resilient" groups also experienced a decrease in the number of

1 average daily steps at three weeks after surgery, but these readily recovered to pre-surgical levels at six  
2 weeks after surgery. These findings show the importance of exploring different physical activity  
3 indicators after surgery as recovery patterns may differ depending on the intensity level or indicator  
4 chosen.

5

### 6 **Pre-surgical risk and resilience predictors of recovery trajectories**

7 Although previous work is inconclusive about the predictive value of biomedical variables for  
8 postsurgical pain recovery after major surgery <sup>1,5,60</sup>, scoliosis severity (i.e., Cobb angle) was predictive  
9 of pain recovery in this study. Unlike previous work<sup>1</sup>, we observed that more severe curvatures were  
10 associated with more favorable pain recovery trajectories. Potentially, another (unmeasured) variable  
11 may have confounded the positive association between the Cobb angle and membership to these more  
12 adaptive pain recovery trajectories.

13 We found that higher *pre-surgical pain intensity* levels were predictive of delayed recovery in  
14 HRQOL after surgery, yet these did not predict postsurgical physical activity levels. Although pre-  
15 surgical pain intensity could not be examined as a predictor of the different postsurgical pain trajectories  
16 (because it was included as the starting point), it was observed that the delayed pain recovery groups  
17 reported higher pain intensity levels before surgery compared to the more favorable pain recovery  
18 groups. This supports findings from earlier work showing that pre-surgical pain intensity is an important  
19 risk factor for worse recovery in pain <sup>1,5</sup> and HRQOL <sup>7</sup> after major surgery in adolescents. In addition,  
20 *catastrophizing about pain before surgery* was found as another risk factor for delayed recovery in pain  
21 intensity, HRQOL, and total number of steps following surgery. Although previous pediatric studies  
22 showed no consistent relation between pre-surgical pain catastrophizing and pain recovery after major  
23 surgery <sup>5,11</sup>, our findings correspond with adult studies and theoretical models (e.g. Fear-Avoidance  
24 Model of pain <sup>66,67</sup>) which identified pain catastrophizing as an important predictor of chronic  
25 (postsurgical) pain. Besides, our study was one of the first to show that pre-surgical pain catastrophizing  
26 levels also predict less favorable recovery in other outcomes such as HRQOL and daily amount of steps  
27 in adolescents after surgery.

1           Furthermore, as expected, higher levels of *psychological flexibility prior to surgery* and  
2 *acceptance of postsurgical pain* at three weeks after surgery were predictive of more favorable recovery  
3 in HRQOL. Higher levels of psychological flexibility were also predictive of more favorable pain  
4 recovery. Unexpectedly, psychological flexibility and acceptance of postsurgical pain did not predict  
5 recovery in physical activity following surgery. As a response to the predominant focus of pain research  
6 on identifying and targeting risk factors for worse pain-related functioning<sup>17</sup>, a growing number of  
7 studies have put forward psychological flexibility and pain acceptance as potential resilience factors that  
8 predict adaptive functioning in the presence of pain<sup>24,25,44,47,68,69</sup>. Importantly, the predominant goal of  
9 promoting psychological flexibility and pain acceptance is maintaining a valued life despite the pain or  
10 physical complaints and *not* reduction in these physical outcomes<sup>20</sup>. Our findings are in line with this  
11 idea showing its predictive value for (more favorable) HRQOL but also for pain outcomes following  
12 surgery. Lower experienced pain intensity levels might be a secondary consequence of generally  
13 improved well-being, although this is a post-hoc speculation which needs further examination.

14

### 15 **Strengths and Limitations**

16           The current study had several strengths. In contrast to previous work recovery trajectories in  
17 functional outcomes and physical activity levels were examined in addition to pain outcomes. It was  
18 also one of the first studies in this domain to *objectively* assess physical activity trajectories up to six  
19 months after surgery (also see<sup>13</sup>). Furthermore, and most importantly, it was the first to focus on  
20 potential resilience factors and demonstrate its adaptive effects of in a surgical context.,

21           Despite these strengths, our findings should be interpreted in the light of several limitations. First,  
22 our sample size was quite small. It may be that in larger samples more meaningful classes with a  
23 substantial number of adolescents could be identified. Although we applied the “common practice” cut-  
24 off value of 5% in order to evaluate if a class was meaningfulness, some classes only consist of 10 or  
25 less people. These classes are useful to provide a first look on potential recovery patterns but first require  
26 replication in larger samples before further interpretations can be made. Moreover, due to this limited  
27 sample size, performance of more complex trajectory analyses was restricted. One reason for the  
28 relatively small sample size might be that we included a very specific clinical population. Over the

1 course of one year, only a limited number of adolescents with AIS undergo scoliosis surgery because  
2 this is often a final treatment option for the more severe cases. Furthermore, adolescents with  
3 comorbidities were excluded. We have tried to meet this recruitment challenge by maximizing our reach  
4 potential through the involvement of the four main (university) hospitals in Flanders who perform  
5 scoliosis surgery. Furthermore, an average of 21% of data were missing across variables and time points.  
6 This limitation may be mitigated by the fact that drop-out is a common challenge in longitudinal studies  
7 with attrition rates ranging between 30 – 70% in other studies with a longitudinal design <sup>70</sup>, the fact that  
8 data was not systematically missing, and that we used Robust Full Information Maximum Likelihood  
9 estimation to handling this missingness. Another limitation was that other potentially important  
10 predictors of postsurgical functioning were not examined because this would lead to too complex models  
11 to estimate given the small sample size. Next, we were restricted by the availability of validated  
12 measures to assess the variables of interest in this context of (sub-)acute postsurgical pain. The GCPS  
13 assesses current pain levels but also requires a recall of average and worst pain over a period of 3 weeks,  
14 which may be problematic for youth experiencing (sub-)acute pain. Also, the CPAQ was originally  
15 developed to assess acceptance of chronic pain and we were one of the first to use this in a subacute  
16 pain context. Furthermore, this study may have been exposed to some threats to generalizability of the  
17 results to the entire population of youth undergoing major surgery. The sample mainly consisted of girls  
18 and was focused on one surgical procedure type. As such, our findings may not generalize to boys  
19 undergoing spinal fusion surgery or to recovery from other surgical procedures. Some level of self-  
20 selection bias may have occurred in that sense that our study may have attracted relatively more resilient  
21 adolescents than those at greater risk for worse recovery (and who might need the most help). Indeed,  
22 about 25% of the patients who were approached for this study declined participation because of the  
23 expected high mental load of participation. Finally, it may be that different trajectories and predictors  
24 may emerge if assessments at later points in time after surgery are included.

25

## 26 **Future directions**

27 Future research could investigate more complex models using bigger samples ( $n > 100$ ). For  
28 instance, it may be examined how recovery in pain, quality of life and physical activity co-vary over

1 time by adding them into one big model. Future work could additionally investigate the role of the social  
2 context. For instance, it has previously been shown that parental behavior may have an important impact  
3 on adolescent recovery after major surgery <sup>5,7,8,71</sup>. Also, the influence of peers and/or health care  
4 providers could be further explored as these may have a critical influence on the functioning of  
5 adolescents who experience (chronic) pain <sup>12,72,73</sup>. Furthermore, there may be other psychosocial  
6 resilience factors that could predict *adaptive* recovery after major surgery. We are aware of one previous  
7 study that found effective pain coping to be such a protective factor for recovery after spinal fusion  
8 surgery <sup>2</sup>, but there remains a wide range of other factors to be explored <sup>17</sup>. Finally, future work could  
9 consider to use alternative pain measures that assess (sub-acute) pain levels, validated direct measures  
10 of psychological flexibility (which is currently not available, to the best of our knowledge), and further  
11 explore the ability to assess acceptance of (sub-)acute pain. In general, replication of our findings in  
12 larger samples is required before making further conclusions about their predictive value for post-  
13 surgical pain recovery in adolescents undergoing major surgery.

14

### 15 **Clinical Implications**

16 If replicated, our findings may inform clinicians about the importance of a multidimensional  
17 recovery assessment (including pain, quality of life, and physical activity levels) after major surgery in  
18 children and adolescents, and to monitor changes in each domain for at least six months after surgery  
19 (and potentially longer; see <sup>4,5</sup>). Our findings also suggest that a screening of psychological risk and  
20 resilience factors should preferably be conducted before surgery in order to target them, prevent worse  
21 outcomes, and stimulate adaptive recovery. This need for pre-surgical psychosocial interventions has  
22 already been expressed by parents of children undergoing surgery <sup>12</sup>. It may be suggested that such  
23 interventions could be conducted in adolescents who are at risk for maladaptive outcomes and could  
24 involve aspects of pain acceptance and psychological flexibility. These concepts are central to  
25 Acceptance and Commitment Therapy (ACT), a cognitive-behavioral therapy that enhances flexibility  
26 in dealing with negative or unwanted experiences and continued engagement in valued activities despite  
27 this adversity. One study in adults undergoing orthopedic surgery gave preliminary evidence for the  
28 efficacy of a one-day pre-surgical ACT intervention in preventing worse recovery <sup>74</sup>. Future research



1 could examine the effectiveness of a brief ACT intervention to promote adaptive functioning in  
2 adolescents as well. Finally, our study supports that accelerometers are a useful tool to objectively assess  
3 physical activity in adolescents who undergo major surgery (see also <sup>13,64</sup>). Although adolescents in this  
4 study did not meet the recommended physical activity levels, achieving this is associated with beneficial  
5 social, mental, and physical health outcomes <sup>63</sup>. Health care providers could use accelerometers to  
6 stimulate physical activity before and after surgery <sup>75</sup>.

7

## 8 **Conclusion**

9 This study identified distinct postsurgical pain, HRQOL, and physical activity trajectories in a  
10 sample of adolescents with AIS undergoing spinal fusion surgery. Pre-surgical pain intensity and pain  
11 catastrophizing increased the risk for lower HRQOL, while pain catastrophizing additionally predicted  
12 lower levels of pain intensity and average daily number of steps after surgery. Psychological flexibility  
13 before surgery and acceptance of postsurgical pain were predictive of adaptive HRQOL outcomes, and  
14 psychological flexibility additionally predicted more favorable recovery in pain after surgery. Average  
15 daily MVPA was not predicted by any of these risk or resilience predictors. Future work could explore  
16 additional risk and resilience predictors of both pain and functional outcomes after major surgery. These  
17 findings could inform targeted screening and intervention prior to surgery to prevent poor and foster  
18 adaptive recovery in adolescents.

19

20

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## 1 TABLES AND FIGURES

2

3 **Table 1.** Sample characteristics.

4 **Table 2.** Parameters of fit for latent class growth analysis for recovery classes in all outcomes

5 **Table 3.** Univariate main effects and pairwise comparison of means to examine relations between class  
6 membership and hypothesized risk and resilience variables

7 **Figure 1.** Flow-chart of recruitment, drop-in, and drop-out of participants during the study.

8 *Note.* Participants were given the opportunity to start their participation after surgery (T1), these are referred to  
9 as “drop-in (new)” participants. Participants were also given the opportunity to re-start their participation after  
10 missing one or more measurement moments, these are referred to as “drop-back-in (old)” participants.

11 **Figure 2.** Plot of fit indices for all outcomes.

12 *Note.* BIC = Bayesian Information Criterion; AIC = Akaike Information Criterion; aBIC = adjusted Bayesian Information  
13 Criterion.

14 **Figure 3.** Plots of final recovery trajectories.

15 *Note.* MVPA= Moderate-to-Vigorous Physical Activity; T0 = 3-1 weeks before surgery;

16 S = Spinal Fusion Surgery; T1 = 3 weeks after surgery; T2 = 6 weeks after surgery; T3 = 6 months after surgery

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**Table 1.** Sample characteristics.

<b>Characteristic</b>	<b>M (<math>\pm</math> SD) or N</b>
<b>Age</b>	15.19 ( $\pm$ 1.56)
<b>Sex</b>	
Female	77
Male	23
<b>Race</b>	
Caucasian	99
Asian	1
<b>Educational level</b>	
Primary School	2
High School	98
<b>Medical variables</b>	
Cobb angle (degrees)	52.19 ( $\pm$ 10.65)
Height (cm)	165.88 ( $\pm$ 8.04)
Weight (kg)	55.05 ( $\pm$ 10.33)
Length of hospital stay	
< 7 days	53
7 – 14 days	31
15 – 21 days	2
<i>Missing info</i>	14

Table 2. Parameters of fit for latent class growth analysis for recovery classes in all outcomes

Outcome	Class	AIC	BIC	aBIC	Entropy	VLMR-LRT	BLRT
						<i>p</i>	<i>p</i>
Pain	1	1428.00	1443.39	1424.45	-	-	-
	2	1330.69	1353.77	1325.36	.90	.001	.000
	3	1308.16	1338.94	1301.05	.88	.347	.000
	4	<b>1276.48</b>	<b>1314.94</b>	<b>1267.58</b>	<b>.92</b>	<b>.021</b>	<b>.000</b>
	5	1264.84	1311.00	1254.17	.90	.178	.000
	6*	1262.19	1316.04	1249.74	.90	.738	.207
	4*	<b>1137.75</b>	<b>1186.37</b>	<b>1123.28</b>	<b>.95</b>	<b>.031</b>	<b>.000</b>
HRQOL	1	2867.14	2882.59	2863.64	-	-	-
	2	<b>2779.16</b>	<b>2802.33</b>	<b>2773.91</b>	<b>.81</b>	<b>.000</b>	<b>.000</b>
	3	2765.08	2795.98	2758.09	.81	.206	.000
	4	2752.25	2790.87	2743.50	.76	.234	.000
	5*	2756.64	2802.98	2746.14	.86	.279	.667
MVPA	1	1584.90	1597.28	1581.51	-	-	-
	2	<b>1564.99</b>	<b>1584.81</b>	<b>1559.57</b>	<b>.72</b>	<b>.141</b>	<b>.000</b>
	3	1554.20	1581.45	1546.74	.81	.236	.000
	4*	1553.15	1587.83	1543.65	.85	.077	.113
Steps	1	1088.29	1100.67	1084.90	-	-	-
	2	1088.98	1108.80	1083.56	.79	.113	.667
	3	<b>1085.87</b>	<b>1113.12</b>	<b>1078.41</b>	<b>.62</b>	<b>.433</b>	<b>.050</b>
	3 <sup>b</sup>	963.93	998.70	951.42	.66	.195	.078
	4*	1087.69	1122.38	1078.20	.70	.682	.667

Notes: bold numbers indicate the output of the final class solution; \* final class solution controlling for Cobb angle at baseline; <sup>b</sup> controlling for age at baseline; AIC = Akaike information criterion; BIC = Bayesian information criterion; aBIC = adjusted BIC; VLMR-LRT = Vuong-Lo-Mendell-Rubin likelihood test; BLRT = Bootstrap

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likelihood ratio test; HRQOL = Health-Related Quality Of Life; MVPA = Moderate-to-Vigorous Physical Activity. \*Model specification was terminated when both VLM-LRT and BLRT were insignificant.

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Table 3. Univariate main effects to examine relations between class membership and hypothesized risk and resilience variables

		Pain (4-class)					
	class	M (SD)	F	df (hyp)	df (error)	p	partial $\eta^2$
Pain Catastrophizing (T0)	1	25.29 (3.92) <sup>a</sup>	3.50	3	376.93	.021* <sup>†</sup>	.15
	2	14.92 (1.73) <sup>b</sup>					
	3	19.57 (2.77) <sup>ab</sup>					
	4	26.00 (4.24) <sup>a</sup>					
Psychological Flexibility (T0)	1	45.86 (3.80) <sup>ab</sup>	1086.23	3	362.07	.019*	.15
	2	52.31 (1.68) <sup>a</sup>					
	3	44.71 (2.69) <sup>b</sup>					
	4	41.17 (4.11) <sup>b</sup>					
Pain Acceptance (T1)	1	41.57 (3.76) <sup>ab</sup>	601.79	3	200.60	.119	.09
	2	46.14 (1.66) <sup>a</sup>					
	3	40.86 (2.66) <sup>ab</sup>					
	4	37.33 (4.06) <sup>b</sup>					
		HRQOL (2-class)					
	class	M (SD)	F	df (hyp)	df (error)	p	partial $\eta^2$
Pain Intensity (T0)	1	5.17 (2.29)	18.55	1	120.22	.000***	.21
	2	2.49 (2.69)					
Pain Catastrophizing (T0)	1	22.70 (10.31)	9.94	1	1119.87	.002**	.13
	2	14.52 (10.80)					
Psychological Flexibility (T0)	1	42.00 (8.93)	22.66	1	2041.39	.000***	.25
	2	53.05 (9.81)					
Pain Acceptance (T1)	1	40.73 (10.59)	5.19	1	530.29	.026* <sup>†</sup>	.07
	2	46.00 (9.80)					
		MVPA (2-class)					
	class	M (SD)	F	df (hyp)	df (error)	p	partial $\eta^2$
Pain Intensity (T0)	1	2.73 (3.47)	1.20	1	9.72	.277	.02
	2	3.80 (2.73)					
Pain Catastrophizing (T0)	1	13.30 (10.92)	2.14	1	279.14	.148	.03
	2	19.04 (11.48)					
Psychological Flexibility (T0)	1	52.30 (10.85)	1.06	1	129.53	.308	.02
	2	48.39 (11.11)					
Pain Acceptance (T1)	1	43.80 (9.21)	0.03	1	2.91	.872	.00
	2	43.21 (10.73)					
		Steps (3-class)					
	class	M (SD)	F	df (hyp)	df (error)	p	partial $\eta^2$
Pain Intensity (T0)	1	4.60 (0.88)	2.22	2	34.68	.117	.07
	2	2.77 (0.55)					
	3	4.08 (0.51)					
Pain Catastrophizing (T0)	1	20.00 (3.36) <sup>ab</sup>	6.56	2	1483.65	.003**	.17
	2	12.39 (2.09) <sup>a</sup>					
	3	22.57 (1.94) <sup>b</sup>					
Psychological Flexibility (T0)	1	50.20 (3.53)	0.49	2	121.27	.617	.02
	2	50.23 (2.19)					
	3	47.50 (2.04)					
Pain Acceptance (T1)	1	44.10 (3.35)	0.06	2	13.53	.942	.00
	2	42.80 (2.08)					
	3	43.47 (1.94)					

<sup>a</sup> Estimated marginal means with a different superscript are significantly different ( $p < .05$ ); \* $p < .05$ , \*\* $p < .01$ , \*\*\* $p < .001$ ; <sup>†</sup> loss of significance when applying Bonferroni correction to account for multiple testing ( $p \leq \alpha/15 = 0.003$ ); Pain classes (1 = Severe-Moderate, 2 = Mild-No, 3 = Moderate-Mild, 4 = Moderate-Severe; HRQOL classes (1 = At-risk, 2 = Resilient; MVPA classes (1 = Low MVPA, 2 = Very Low MVPA), Steps classes (1 =

Resilient Low Steps, 2 = Resilient High Steps, 3 = Declined Steps). HRQOL = Health-Related Quality Of Life, MVPA = Moderate-to-Vigorous Physical Activity. T0 = before surgery; T1 = 3 weeks after surgery

1

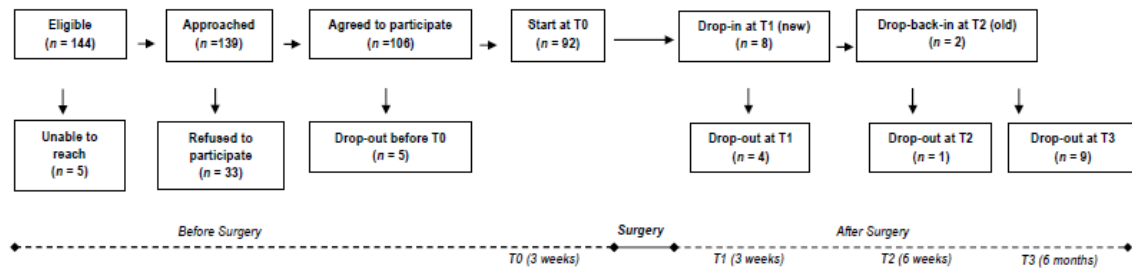


Figure 1. Flow-chart of recruitment, drop-in, and drop-out of participants during the study.  
 Note. Participants were given the opportunity to start their participation after surgery (T1), these are referred to as “drop-in (new)” participants. Participants were also given the opportunity to re-start their participation after missing one or more measurement moments, these are referred to as “drop-back-in (old)” participants.

2

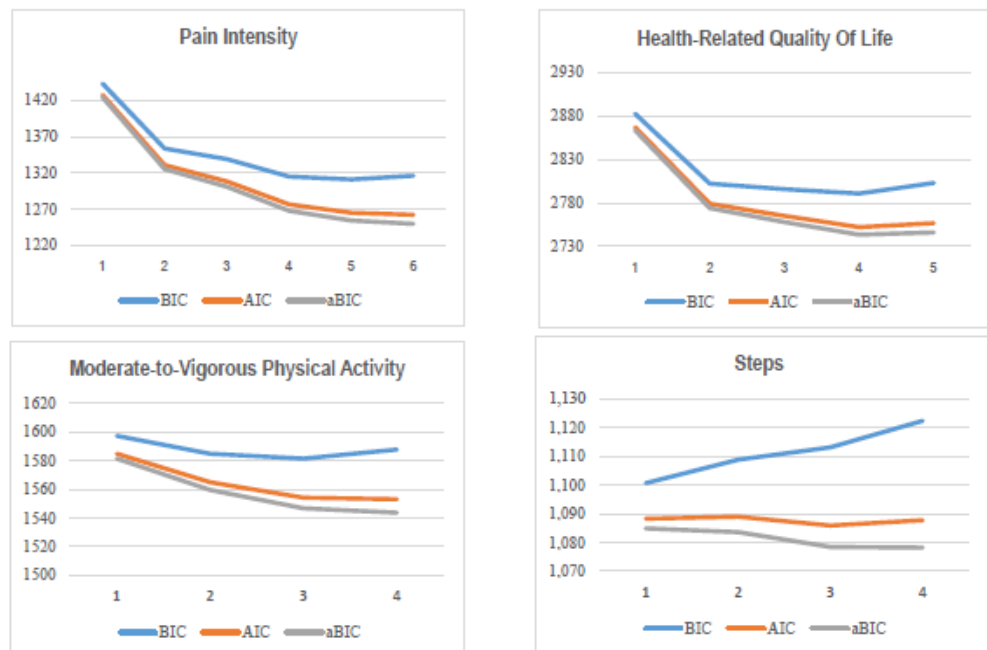
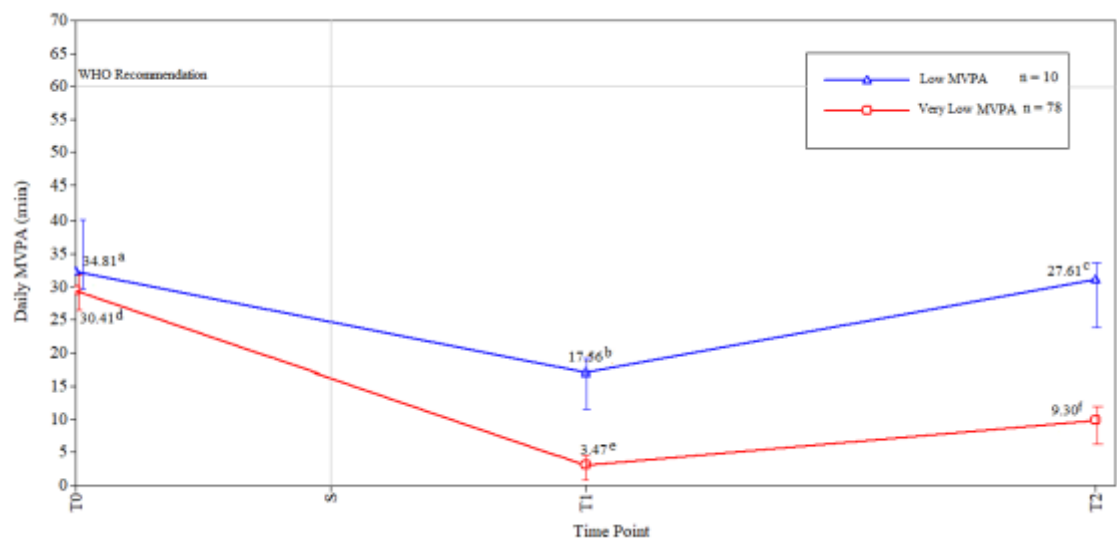
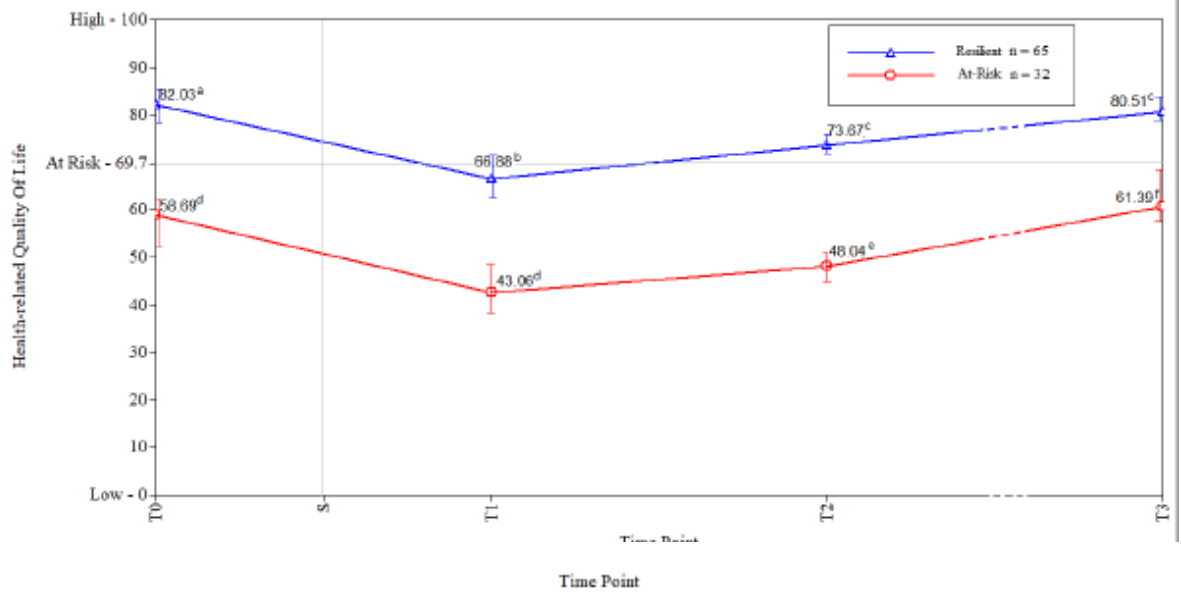
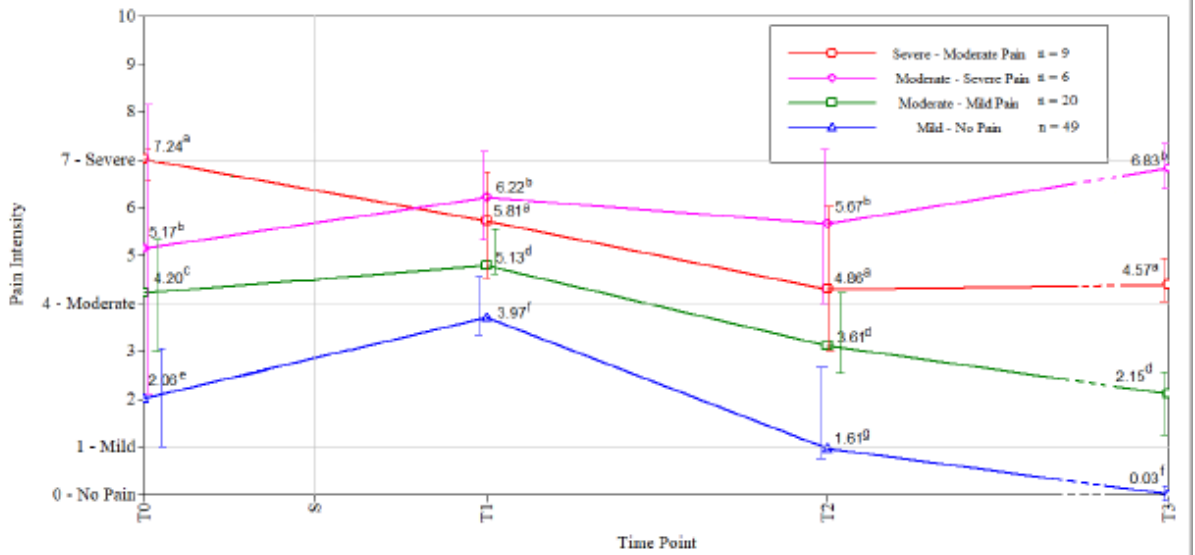


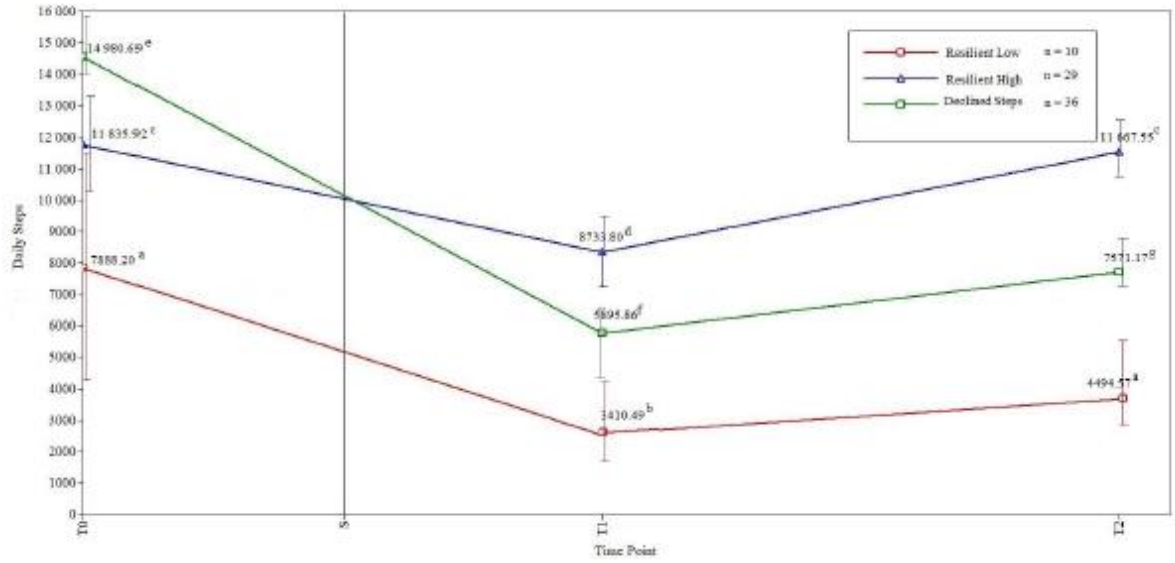
Figure 2. Plot of fit indices for all outcomes.  
 Note. BIC = Bayesian Information Criterion; AIC = Akaike Information Criterion; aBIC = adjusted Bayesian Information Criterion.

3



1

2



1

**Figure 3.** Plots of final recovery trajectories with estimated marginal means and 95% confidence intervals.  
 Note. MVPA= Moderate-to-Vigorous Physical Activity; T0 = 3-1 weeks before surgery; S = Spinal Fusion Surgery; T1 = 3 weeks after surgery; T2 = 6 weeks after surgery; T3 = 6 months after surgery

2