

Provisioning tactics of great tits in response to long-term brood manipulations differ across years

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1 **Running head:** Test of variance-sensitive provisioning

2 **Provisioning tactics of great tits (*Parus major*) in response to long-term brood size**
3 **manipulations differ across years**

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25 **Authorship statement**

26 KJM, AM, A-LO, MN, JW, BK, and ND designed the experiment. KJM, AM, A-LO, MN, YAA,
27 and ND performed the field work. AM and ALO coordinated the video analyses and compiled
28 the database. KJM, YAA and NJD performed the statistical analyses with input from all
29 coauthors. KJM and DFW wrote the manuscript. All coauthors contributed to revisions.

30

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51 **Data accessibility**

52 All data used in the analyses presented in this paper will be deposited on Dryad upon publication
53 of the paper.

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4 1 **Lay summary:** Parents adjust their provisioning behaviour in response to offspring
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6 2 demand, but the way they do this depends on ecological conditions. Parent great tits (*Parus*
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8 3 *major*) responded to experimentally increased brood demand by working harder and
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10 4 decreasing prey selectivity in a year with favourable ecological conditions. In a more
11
12 5 challenging year, parents did not meet increased offspring demand by shifts in average
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14 6 behaviour, and instead exhibited shifts in the variance in their behaviour.
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20 7 **Title:** Provisioning tactics of great tits (*Parus major*) in response to long-term brood size
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22 8 manipulations differ across years
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26 9 **Abstract:** Parents provisioning their offspring can adopt different tactics to meet increases
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28 10 in offspring demand. In this study, we experimentally manipulated brood demand in free
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30 11 living great tits (*Parus major*) via brood size manipulations and compared the tactics
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32 12 adopted by parents in two successive years (2010 and 2011) with very different ecological
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34 13 conditions. In 2011, temperatures were warmer, there were fewer days with precipitation,
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36 14 and caterpillars (the preferred prey of great tits) made up a significantly larger proportion of
37
38 15 the diet. In this 'good' year, parents responded to experimental increases in brood demand
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40 16 by decreasing mean inter-visit intervals (IVIs) and reducing prey selectivity, which
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42 17 produced equal average long-term delivery of food to nestlings across the brood size
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44 18 treatments. In 2010, there was no evidence for effects of brood size manipulations on mean
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46 19 IVIs or prey selectivity. Consequently, nestlings from enlarged broods experienced
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48 20 significantly lower long-term average delivery rates compared with nestlings from reduced
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50 21 broods. In this 'bad' year, parents also exhibited changes in the variance in inter-visit
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4 22 intervals (IVIs) as a function of treatment that were consistent with variance-sensitive
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7 23 foraging theory: variance in IVIs tended to be lowest for reduced broods and highest for
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10 24 enlarged broods. Importantly, this pattern differed significantly from that observed in the
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12 25 ‘good’ year. We therefore found some support for variance-sensitive provisioning in the
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14 26 year with more challenging ecological conditions. Taken together, our results show that
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16 27 variation in brood demand can result in markedly different parental foraging tactics
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19 28 depending on ecological conditions.

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22 29 **Keywords:** brood demand, brood size manipulation, heterogeneous residual variance,
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25 30 *Parus major*, provisioning behaviour, variance-sensitivity, risk-sensitivity
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31 Introduction

32 In many organisms, parents care for offspring and they adjust this care in response to cues
33 that reflect the benefits and/or the costs of parental care (e.g. Clutton-Brock, 1991; Leonard
34 and Horn, 1998; Nakagawa and Schielzeth, 2017; Royle et al., 2012; Westneat et al., 2011;
35 Williams, 1966; Winkler, 1987; Wright et al., 1998; Wright and Leonard, 2010). In birds,
36 parents provisioning young in the nest can employ a variety of tactics to cope with
37 increases in brood demand (Wright et al., 1998; Ydenberg, 2008). Parents can increase the
38 time and/or energy allocated to the current brood at the cost of their own expected future
39 fitness (Moreno et al., 1995; Ydenberg, 2008). For example, they might switch from a net
40 energy maximization (i.e. the difference between energy provisioned to young and energy
41 expended by the parent) to a gross energy maximization (i.e. total energy provisioned to
42 young, ignoring costs associated with provisioning effort) (Lifjeld, 1989). The decision to
43 increase provisioning can show different forms, because parents can reduce the time
44 allocated to other activities (e.g. nest defense, Markman et al., 1995; self-feeding, Martins
45 and Wright, 1993) and in this way they can increase their provisioning rate and/or they can
46 bring larger or higher quality items per delivery (e.g. Wright, 1998) (Figure 1a).

47 Parents may also shift the relative nutritional quality versus energetic quantity of food
48 provided. Prey items that provide the best energetic returns may not be the same as those
49 containing the optimal combination of nutritional compounds for offspring growth (see
50 Wright et al., 1998 and references therein). If so, parents working to meet increased brood
51 demand may reduce prey selectivity, and increase the rate of energy delivery to the nest at a

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4 52 cost of reduced nutritional value via increased frequency of visits to the nest and/or
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7 53 increased energy delivered per visit (Figure 1b).
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10 54 Though less commonly appreciated, parents can also adopt strategic shifts in their use of
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12 55 provisioning options that are associated with higher variance in prey delivery (i.e. variance-
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14 56 sensitivity, a.k.a. risk-sensitivity) as a way of mitigating increased brood demand (Westneat
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17 57 et al., 2013; Ydenberg, 1994, 2008). Variance in prey delivery generates an important
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19 58 selective force when it causes deviations in energy delivery above and below the expected
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21 59 mean delivery, because these deviations can have asymmetrical effects on offspring fitness
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23 60 (Figure 2). When offspring are in a low energy state or on a negative energy budget, they
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25 61 likely experience a convex (accelerating) utility function (left side in Figure 2). If so, they
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27 62 will benefit disproportionately when variance in intake is increased because of the
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29 63 disproportionate impact of higher than expected delivery on fitness (i.e., mean fitness is
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31 64 higher than the fitness of the mean intake). Parents would thus maximize their expected
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33 65 fitness return by preferring foraging options that increase variance in return. In contrast,
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35 66 when offspring are in a high energy state or on a positive energy budget, they likely
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37 67 experience diminishing returns with a concave (decelerating) utility function (right side in
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39 68 Figure 2). They then will benefit disproportionately from reduced variance in intake (i.e.,
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41 69 fitness from the mean intake is higher than the mean fitness from the distribution of
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43 70 intakes). Parents are said to be “variance sensitive” if they use variance as a cue when
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45 71 making provisioning decisions (Ydenberg, 2008). Optimality theory predicts that parental
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47 72 variance-sensitivity will be adaptive, because changes in variance in energy delivery
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49 73 according to the state of the offspring increase offspring fitness. When faced with offspring
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4 74 that are experiencing a convex utility function (e.g. offspring in poor condition), parents
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6 75 should be "variance-prone" and seek foraging options that are more variable and variance in
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9 76 delivery should increase. Conversely, when faced with offspring experiencing a concave
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11 77 utility function (e.g. offspring in good condition), parents should be "variance-averse" and
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13 78 seek less variable foraging options, and variance in delivery to the nest should decrease
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16 79 (Figure 1c).
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20 80 Shifts in variance in energy delivery to offspring could come about in several ways that
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22 81 depend upon the underlying prey distributions (i.e., variance in encounter rates with
23
24 82 particular prey types) (Sutherland and Anderson, 1987). For example, foragers can alter the
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26 83 variance in energy gain by strategically altering their probability of rejecting or accepting
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28 84 lower ranking prey types (Weissburg, 1991). Variance in prey encounter rates may also
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30 85 differ across foraging patches (Sutherland and Anderson, 1987), and foragers may be able
31
32 86 to manipulate variance in prey encounter rates by choosing locations with more or less
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34 87 variability or by altering patch departure times (Stephens and Charnov, 1982). All of these
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36 88 would alter the variance in returns per trip that are experienced by offspring.
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41 89 Assessing variance sensitivity and the specific ecological conditions by which it comes
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43 90 about requires manipulations of offspring demand on parents (Ydenberg, 2008). Brood size
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45 91 manipulations are the most commonly used method for manipulating brood demand
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47 92 (Ydenberg, 2008). For adults provisioning young, having more young logically increases
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49 93 the required energy because, all else being equal (i.e. food availability, energetic costs
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51 94 imposed by weather conditions, etc.), more young require more energy. The extent to which
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53 95 the increased energy demands of enlarged broods translate to increased probability of
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4 96 experiencing meaningful energetic shortfalls will depend on ecological conditions (e.g.
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7 97 prey availability and/or ambient conditions that influence nestling energy expenditure;
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9 98 Moore, 2002). When enlarged broods do experience an increased probability of energetic
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11 99 shortfall, parents should generally increase their relative preference for prey with more
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14 100 variable encounter rates (Moore, 2002; Ydenberg, 2008) (see Figure 1c and Figure 2).

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17 101 Here we report the results of a two-year study during which we altered offspring demand on
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19 102 parent great tits (*Parus major*) using brood size manipulations. We estimated changes in
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21 103 mean inter-visit intervals (IVIs), mean load size, mean short-term delivery (load size/time
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23 104 for each provisioning visit), and the relative contribution of caterpillars (the preferred prey
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25 105 of great tits, Naef-Daenzer et al., 2000; Wilkin et al., 2009) to the diet. Additionally, we
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27 106 quantify variances in IVIs, load size and short-term delivery experienced by nestlings. We
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29 107 also estimated the effect of experimentally manipulated brood demand on measures of
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31 108 long-term provisioning: provisioning rate (visits per parent per 30 minute observation) and
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33 109 long-term delivery (sum of load sizes per 30 minute). We evaluated the extent to which
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35 110 changes in means and variances in provisioning behaviour followed predictions from the
36
37 111 three strategic options outlined above for coping with increased brood demand: increase
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39 112 provisioning effort, shift in energetic quantity relative to nutritional quality, or variance
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41 113 sensitivity (Figure 1). A fortuitous difference in ecological conditions (detailed in Nicolaus
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43 114 et al., 2015) between the two years allowed us to compare the strategies adopted by parents
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49 115 across the two study years in light of these contrasting ecological conditions.
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117 **Materials and Methods**118 *Study population*

119 This study was carried out in a nestbox population of great tits (*Parus major*) located in
120 southwestern Germany. The study site consists of 12 plots covering a total area of
121 approximately 120 ha. Each plot is comprised of 50 nestboxes arranged in a regular grid
122 (50 m between adjacent nestboxes). All nestboxes were monitored throughout the breeding
123 season (*circa* March to August) by visiting them at least twice per week in order to record
124 lay date (date of first egg, back calculated assuming 1 egg laid per day), clutch size, onset
125 of incubation, hatching and fledging date and the number of young fledged (see Nicolaus et
126 al., 2015 for more details).

127 When nestlings were 3 days old, we performed brood size manipulations (BSMs) such that
128 parents received either a reduced brood (net change of -3 nestlings), a control brood (net
129 change of 0 nestlings, but nestlings still swapped between nests), or an enlarged brood (net
130 change of +3 nestlings). Nestlings were exchanged between nests of the same age and with
131 a similar average nestling body mass. After the manipulation, half of the nestlings in all
132 manipulated nests came from another nest. Nests for which there was no suitable matching
133 brood were left unmanipulated and are not discussed further here (see Nicolaus et al., 2015
134 for further details on the BSMs). BSMs were performed blind with respect to parental traits
135 (e.g., behaviour, morphology) that may in turn have affected their response to the BSM.

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4 136 Furthermore, BSMs were stratified across natural brood sizes, so there were no BSM-
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7 137 related differences in pre-manipulation brood size.
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10 138 *Provisioning behaviour*
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13 139 Provisioning behaviour was monitored at a total of 83 nestboxes that received BSMs (N =
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15 140 42 in 2010 and N = 41 in 2011). This represents less than 20% of all first broods in each
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17 141 year (N = 223 first broods in 2010, N = 216 first broods in 2011). To record provisioning
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19 142 behaviour, small infrared cameras (CDD Bird Box Camera with IR Night Vision 420TV
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21 143 lines) were installed in nestboxes when nestlings were 10 days old by exchanging the side
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23 144 door of the nestbox with a small wooden box containing the camera. Nests were raised
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25 145 circa 2 cm by placing a piece of foam beneath the nest material to ensure that the entire nest
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27 146 cup could be filmed. Additionally, a clear piece of Plexiglas was placed between the nest
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29 147 and the side compartment that contained the infrared camera to prevent parents from sitting
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31 148 in the side compartment (Supplementary Figure S1). The following day (nestling day 11),
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33 149 nestlings were weighed and given unique markings on the top of their heads with acrylic
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35 150 paint (as part of another experiment).
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42 151 On nestling day 12, at approximately 8h00, a portable recording device (Archos 5 Internet
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44 152 Media Tablet) and power supply were connected to the nestbox camera using 20 m long
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46 153 cables. This allowed the cameras to be turned on to start recordings at a distance from the
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48 154 nestbox. Observations began 30 minutes after camera installation. For observation period 1,
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50 155 the observer walked up to the nestbox (within 1 m) before starting the recording device.
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53 156 The observer then left the immediate vicinity of the nestbox (> 100 m) for 30 minutes (the
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4 157 duration of observation period 1) before returning to interrupt and resume the recording (at
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7 158 20 m distance from the nestbox) for another 30 minutes (observation period 2). The
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9 159 observer again left the vicinity of the nestbox for the duration of observation period 2. The
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11 160 difference in close (within 1 m) versus far (circa 20 m) approach to the nestbox between
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14 161 observation period 1 and 2 was part of another experiment to investigate potential effects of
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16 162 short human disturbances at the nestboxes on provisioning behaviour (Mutzel et al., *In*
17
18 163 *preparation*). Importantly, Mutzel et al. (*In preparation*) found no evidence of year-specific
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20 164 effects of these two disturbance levels on parental provisioning behaviour, suggesting that
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22 165 any year-specific differences observed in the present analyses as a function of the BSM are
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24 166 not due to year-specific responses to these two disturbance treatments. Following the two
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26 167 observation periods included in the present study, additional provisioning data were
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28 168 collected as part of another experiment on provisioning behaviour under different levels of
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30 169 perceived nest predation threat (Mutzel et al., *In preparation*).
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36 170 For each parental visit to the nestbox, the following data were extracted from the videos:
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38 171 bird identity (male or female parent, determined from plumage characteristics), the time (to
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40 172 the nearest second) that a bird entered the nestbox, the load size relative to parent's bill
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42 173 volume (e.g. 1 = volume of bill, 2 = twice the volume of the bill, scored to the nearest 0.1
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44 174 bill volumes) and the time the individual left the nest (to the nearest second). Prey items
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46 175 were classified as preferred (i.e. caterpillar), or non-preferred (Naef-Daenzer et al., 2000;
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48 176 Wilkin et al., 2009). The combined effects of load size and prey type were used to assess
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50 177 evidence for changes in energetic value of prey delivered to the nest (e.g. Wright et al.,
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52 178 1998). For the data used in the present study, 5 observers recorded the timing of nest visits,
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4 179 and 2 observers recorded prey type and load size, from the videos. The two observers for
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6 180 prey type and load sizes were trained until among-observer consistency was high (Pearson's
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8 181 product moment correlation = 0.90, N = 173).

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12 182 From these data, we calculated inter-visit intervals (IVIs) and inter-feed intervals (IFIs).
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14 183 IVIs represent the interval between successive provisioning visits to the nest by the same
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16 184 parent, and were calculated as (nestbox entry time for visit N) – (nestbox entry time for
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18 185 visit N-1). IFIs represent the interval between successive provisioning visits, regardless of
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20 186 the identity of the parent. Thus, short-term delivery to the brood was estimated for each
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22 187 provisioning visit as load size/IFI.

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27 188 We also estimated the combined effects of changes in means and variances in short-term
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29 189 provisioning decisions (IVIs and load sizes) on long-term provisioning rates (sum of visits
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31 190 by each parent over the total observation time), and long-term nest level delivery (sum of
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33 191 load sizes by either parent over the total observation time) (see Bateson and Kacelnik, 1996
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35 192 for discussions of different rate currencies).

36 37 38 39 40 193 *Statistical analyses*

41 42 43 194 *Brood size manipulations*

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46 195 To verify that the brood size manipulations (carried out on nestling day 3) resulted in
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48 196 differences in brood size on the day that provisioning behaviour was filmed (nestling day
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50 197 12), we used a mixed effects model with natural clutch size (within-year centered), and the
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52 198 unique combination of year and treatment (6 levels: 2010R, 2010C, 2010E, 2011R, 2011C,
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54 199 2011E) as fixed effects. Plot (12 levels, 11 of which were present in both years) was
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4 200 included as a random effect. We did not have sufficient replication of nestboxes (only 2
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6 201 nestboxes were replicated across years) or parent IDs (only 4 parent ids replicated across
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8 202 years) to meaningfully include these as random effects. We used square-root transformed
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10 203 brood size in our analyses (to approach a normal distribution) and modelled these
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12 204 transformed data with Gaussian errors. Analyses were performed using data from all BSM
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14 205 nests for which we observed provisioning behaviour (2010: N = 15 reduced, N = 13
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16 206 control, N = 14 enlarged; 2011: N = 14 reduced, N = 13 control, N = 14 enlarged). We
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18 207 tested for a year \times treatment interaction by comparing the effect of treatment (difference in
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20 208 brood size between reduced and enlarged broods) across the two study years. These
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22 209 analyses revealed year-specific effects of the BSM on realized brood sizes on the day of
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24 210 provisioning observations (see results). The two years were therefore treated separately in
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26 211 all subsequent analyses by modelling provisioning behaviour as a function of year-specific
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28 212 brood size manipulations (i.e. 6-level factor: 2010R, 2010C, 2010E, 2011R, 2011C,
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30 213 2011E).

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38 214 *Provisioning behaviour*

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41 215 *Visit-level decisions*

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44 216 We tested whether there was evidence for differences in means and variances in visit-level
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46 217 provisioning behaviour across the two study years as a function of the BSM. We assess this
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48 218 from the perspective of each parent and of the nestlings. To do this for parent level
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50 219 decisions, we constructed univariate mixed effects models with IVI, load size or prey type
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52 220 (preferred or non-preferred) as the response variable. Prey type was analyzed to test for
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4 221 changes in prey selectivity. Caterpillars typically make up the majority of prey items
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6 222 provisioned to young great tits and are thought to represent the most profitable prey type
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9 223 (energy per unit time) because they are provisioned in higher frequencies than expected
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11 224 based on natural abundances (Naef-Daenzer et al., 2000; Wilkin et al., 2009). Thus, within
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14 225 a given year, a decrease in the proportion of caterpillars with increasing brood demand
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16 226 would be consistent with expansion of the diet to include non-preferred prey types to cope
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18 227 with increased demand. We coded each provisioned prey as being preferred (1 =
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21 228 caterpillars) or non-preferred (0 = others), and constructed a binary mixed effect model
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23 229 with prey type (preferred or non-preferred) as the response variable.
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26 230 Each of these parental visit-level decisions (IVI, load size, and prey type) was modelled as
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28 231 a function of the combination of year and treatment (6 level factor: 2010R, 2010C, 2010E,
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30 232 2011R, 2011C, 2011E) and individual parent (2010: 71 levels, 2011: 79 levels), nestbox
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32 233 (2010: 41 levels, 2011: 40 levels) and block identity (the unique combination of nestbox
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34 234 and observation period: each nestbox had up to two unique observation period identities as
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36 235 described above ; 2010: 78 levels, 2011: 79 levels) were fitted as random effects. The
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38 236 number of levels for nestboxes in 2010 and 2011 are lower than the total number of nests
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40 237 observed because of problems with video recording or insufficient visits (i.e. less than 2
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42 238 visits per parent) to calculate parental IVI at 2 nests in 2010, and 1 nest in 2011. To test
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44 239 whether there were treatment-related differences in the variance in IVI or load size, residual
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46 240 variances were modelled for each year by treatment combination (Cleasby and Nakagawa,
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48 241 2011; Hadfield, 2010; Westneat et al., 2015). Heterogeneous errors cannot be modelled for
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4 242 binary data, therefore, we do not assess BSM-related effects on variances in prey types
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6 243 provisioned.
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10 244 We also assessed the combined effects of parental IVI and load size decisions on means and
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12 245 variances in short-term delivery received by each nestling. Short-term delivery was
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14 246 estimated for each visit regardless of the identity of the parent as load size/IFI (units: bill
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16 247 volumes/second). As with the analyses for parent-level data, the unique combination of year
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18 248 and BSM treatment was included as a fixed effect. We additionally included brood size as a
19
20 249 fixed effect to account for the average effect of brood size on delivery independent of BSM
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22 250 treatment. Note that because delivery was ln-transformed prior to analyses, the estimated
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24 251 brood size effect on delivery is non-linear on the observed scale. A significant Year:BSM
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26 252 effect on delivery over and above the effect of brood size would indicate differences in per-
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28 253 nestling delivery as a function of the manipulations. Nestbox and block identity were
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30 254 included as random effects, and residual variance was modelled for each year by treatment
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32 255 combination. Parent identity was not included as a random effect because the response
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34 256 variable was constructed from the combined effects of both parents at a nest.
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41 257 IVI and short-term delivery were ln-transformed, and load size was square root transformed
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43 258 prior to analyses. All of the above models were then constructed with Gaussian error
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45 259 distributions.

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49 260 *Long-term consequences of visit-level decisions*

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53 261 Variance-sensitive foraging theory predicts changes in preference or aversion to variance in
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55 262 response to changes in the long-term expectation of energy shortfall of nestlings. Because
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5 263 our short-term measures of provisioning were transformed prior to analyses using non-
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7 264 linear transformations, back-transformation cannot be used to translate the mean on the
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9 265 latent scale to the mean on the observed scale (Nakagawa and Schielzeth, 2017). Although
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11 266 methods are available to approximate the mean on the observed scale (e.g. 2nd order Taylor
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13 267 expansions or simulations, see Nakagawa and Schielzeth, 2017), we instead calculated
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16 268 long-term provisioning rates (for each parent individually) directly from the raw data (i.e.,
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18 269 number of visits per parent per 30 minute observation). Along similar lines, extrapolating
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21 270 visit-level rates to infer long-term average rates produces biased estimates, and the
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23 271 magnitude of bias depends on the specific data structure (for discussion and examples, see
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26 272 Bateson and Kacelnik, 1996; Templeton and Lawlor, 1981; Welsh et al., 1988). Therefore,
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28 273 we also calculated long-term average nest-level delivery from the raw data (i.e., sum of
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31 274 load sizes delivered per 30 minute observation for both parents combined).

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34 275 To analyze BSM-related effects on parent level long-term average provisioning effort, we
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36 276 modelled (as above) long-term visit rates per parent (sum of visits by a given parent) as a
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38 277 function of a 6 level factor representing the unique combination of year and treatment
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41 278 (2010R, 2010C, 2010E, 2011R, 2011C, 2011E). Nestbox, Block ID (a unique combination
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43 279 of nestbox and observation period) and parent ID were also included as random effects.

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45 280 Although our dependent variable constituted count data, counts of visits do not arise from a
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48 281 Poisson process because provisioning visits by the same parent are not independent of one
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51 282 another. Consequently, we did not analyse these data using Poisson error distributions, but
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53 283 instead used $\ln(n+1)$ transformed counts of parental visit for analyses with Gaussian error
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55 284 distributions.

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4 285 To analyze BSM-related differences in long-term delivery experienced per nestling, we
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6 286 modelled long-term nest-level delivery (sum of load sizes from both parents per 30 minute
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9 287 observation) as a function of a 6 level factor representing the unique combination of year
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11 288 and treatment (2010R, 2010C, 2010E, 2011R, 2011C, 2011E). We included absolute brood
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14 289 size to assess the average care received per nestling. Thus, significant Year:BSM effects on
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16 290 delivery over and above the effect of brood size would indicate differences in per-nestling
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18 291 delivery as a function of the manipulations. Nestbox and Block ID (a unique combination
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21 292 of nestbox and observation period) were included as random effects. Nestling-level visit
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23 293 rate was normally distributed and therefore was not transformed.
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26 294 All mixed effects models were fitted in the R statistical environment version 3.2.3 (R
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28 295 Development Core Team, 2015). Models to estimate year and BSM specific values for both
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30 296 means and residual variances (IVI, load size and short term delivery) were fitted using
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32
33 297 Monte Carlo Markov chains in the MCMCglmm package (Hadfield, 2010), which retrieves
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35 298 posterior distributions of the estimated parameters. Details of parameter estimates are
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37 299 provided in the electronic supplementary material (see Supplementary Text S1).
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40 300 Multivariate mixed-effects models were also constructed in ASREML to validate the
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42 301 Bayesian analyses by implementing an alternative REML-approach, which corroborated
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44 302 our findings (Results not shown).
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48 303 We did not model heterogeneous residual errors for measures of long-term provisioning
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50 304 (long-term provisioning rate and long-term delivery per nestling) because we had only two
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52 305 measures per individual parent (or per nestbox), and therefore no power to detect
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55 306 heterogeneous residual errors. These models were constructed using the lmer function from
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4 307 the lme4 package in R (Bates et al., 2015). We used the ‘sim’ function of the ‘arm’ package
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6 308 to simulate values of the posterior distribution of the model parameters (Gelman and Su,
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8
9 309 2015).

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11
12 310 We evaluated support for treatment effects based on estimated effects sizes and their 95%
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14 311 credible intervals (Cumming and Finch, 2005; Nakagawa and Cuthill, 2007). This approach
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16
17 312 is advocated to avoid drawing dichotomous conclusions (e.g. accept or reject the null
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19 313 hypothesis) based on data which can show a continuous range of support (or lack of
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21 314 support) for a given interpretation (Cohen, 1990; Wasserstein and Lazar, 2016). As a
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23
24 315 reference for readers less familiar with CIs: a 95% CI is roughly equivalent to a value of p
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26 316 ≤ 0.05 when the null distribution is outside the bounds of the CI (i.e. testing for a difference
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28 317 between two groups, $H_1: \beta \neq 0$, with no *a priori* prediction for whether the difference
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30 318 would be positive or negative), or $p \leq 0.025$ when there is a directional prediction
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32 319 (Cumming and Finch, 2005) (e.g. $H_1: \beta > 0$). We use the terms “strong support for an
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34 320 effect” or “significant” when the 95% CI does not overlap zero, moderate support when the
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36 321 estimated effect is removed from zero but the 95% CI overlaps zero by up to 15%, and no
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38 322 support for an effect (or support for lack of an effect) when the estimate is centered on zero.
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45 46 47 324 **Results**

48 49 50 325 ***Brood size manipulations***

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53 326 In both study years, reduced broods were significantly smaller than enlarged broods on the
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55 327 day of filming (control broods were intermediate to the reduced and enlarged broods)
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4 328 (Table 1, Figure 3). The overall variation in brood size produced via the BSM (i.e.
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6 329 difference in number of nestlings between enlarged and reduced treatments) was greater in
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9 330 2011 ($\beta = 1.03$, 95% CI = 0.82, 1.23) compared with 2010 ($\beta = 0.69$, 95% CI = 0.49, 0.88);
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11 331 the presence of a year-difference was strongly supported (Year \times Treatment interaction: $\beta =$
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13 332 0.34, 95% CI = 0.05, 0.62, Figure 3). There was also an effect of natural clutch size on
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16 333 brood size on the day of filming (Table 1).

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20 334 *Effects of the BSM on mean visit-level provisioning behaviour*

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23 335 On average, mean IVIs were shorter in 2010 compared with 2011 (Table 2, Figure 4a), and
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25 336 there was moderate support for a difference in effect of the BSM across the years ($\beta = 0.26$,
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27 337 95% CI = -0.09, 0.69, Bayesian p-value = 0.09). Post-hoc year-specific analyses support
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29 338 year-related differences in IVIs in response to the BSM. In 2010, there was no support for
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31 339 BSM-related differences in mean IVIs (difference between reduced and enlarged broods: β
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33 340 = 0.05, 95% CI = -0.23, 0.33). In contrast, in 2011, there was strong support for IVIs
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35 341 decreasing across the BSM categories (difference between reduced and enlarged broods: β
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37 342 = 0.32, 95% CI = 0.02, 0.58).

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42 343 Average load size was smaller in 2010 compared with 2011 (Table 2, Figure 4b), and there
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44 344 was also moderate support that the effects of the BSM on load sizes differed across the two
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46 345 years ($\beta = 0.06$, 95% CI = -0.05, 0.20, Bayesian p-value = 0.15). In 2010, there was
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48 346 moderate support that load sizes increased across BSM categories (difference between
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50 347 reduced and enlarged broods: $\beta = 0.07$, 95% CI = -0.01, 0.16, Bayesian p-value = 0.05),
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4 348 but there was no support for any BSM differences in load size in 2011 (difference between
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6 349 reduced and enlarged broods: $\beta = 0.02$, 95% CI = -0.06, 0.11).
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10 350 Caterpillars made up a more substantial portion of the diet in 2011 compared with 2010
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12 351 (Table 2). In 2010, the proportion of caterpillars did not differ as a function of the BSM
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14 352 (difference between reduced and enlarged broods: $\beta = 0.20$, 95% CI = -0.52, 1.08). In
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16
17 353 contrast, in 2011, the proportion of caterpillars was greater in reduced broods compared
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19 354 with both control ($\beta = 1.02$, 95% CI = 0.13, 1.82) and enlarged ($\beta = 0.94$, 95% CI = 0.06,
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21 355 1.63) broods.
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25 356 Short-term delivery rates increased significantly with increasing current brood size ($\beta =$
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27 357 0.12, 95% CI = 0.07, 0.18). After accounting for the effect of brood size, short-term
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29 358 delivery per nestling was similar in both study years and across the BSM treatment (Table
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31 359 2, Figure 4c). There was no evidence for BSM-related differences in mean short-term
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33 360 delivery (load size/IFI, controlling for current brood size) in either 2010 (difference
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35 361 between reduced and enlarged broods: $\beta = -0.22$, 95% CI = -0.46, 0.15) or 2011 (difference
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37 362 between reduced and enlarged broods: $\beta = -0.05$, 95% CI = -0.42, 0.36) implying no BSM-
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39 363 related differences in per nestling delivery.
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45 364 *Effects of the BSM on variance in visit-level provisioning behaviour*

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48 365 We found strong support for an effect of the BSM on variance in IVIs that differed across
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50 366 the two study years (difference between enlarged and reduced treatment between 2010 and
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52 367 2011: $\sigma^2 = 0.25$, 95% CI = 0.02, 0.47) (Table 3, Figure 4d). There was strong support for
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54 368 variances in IVIs decreasing across BSM categories in 2011 ($\sigma^2 = -0.19$, 95% CI = -0.37, -
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4 369 0.01), which is opposite to the pattern predicted by variance-sensitive foraging theory.
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6 370 However, in 2010, BSM-related differences in residual variance in IVIs shifted
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9 371 significantly in the direction predicted by variance-sensitive foraging theory (year x
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11 372 treatment interaction). Reduced broods had the lowest residual variance in IVIs and
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13 373 enlarged broods had the highest residual variance in IVIs, although the difference was
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16 374 relatively weak (difference between reduced and enlarged broods: $\sigma^2 = 0.05$, 95% CI = -
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18 375 0.07, 0.17, Bayesian p-value = 0.22).

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21 376 We found no support for BSM effects on variance in load sizes in either 2010 (difference
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23 377 between reduced and enlarged broods: $\sigma^2 = 0.01$, -0.01, 0.02) or 2011 (difference between
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25 378 reduced and enlarged broods: $\sigma^2 = 0.01$, 95% CI = -0.01, 0.03) (Table 3, Figure 4e). By
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27 379 contrast, we observed year-specific patterns of variance in short-term delivery rates
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29 380 experienced by nestlings as a function of the BSM (Table 3, Figure 4f). In 2010, the
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31 381 variance in prey delivery experienced per nestling did not differ between control and
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33 382 enlarged broods ($\sigma^2 = 0.05$, 95% CI = -0.22, 0.36), but there was strong support for
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35 383 nestlings from reduced broods experiencing lower variance in short term delivery rates
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37 384 compared with nestlings from enlarged broods ($\sigma^2 = -0.23$, 95% CI = -0.45, -0.01) and
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39 385 moderate support for nestlings from reduced broods experiencing lower variance in short-
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41 386 term delivery rates compared with nestlings from control broods ($\sigma^2 = -0.15$, 95% CI = -
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43 387 0.40, 0.04). In contrast, in 2011, variance in short-term delivery did not differ
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45 388 systematically across the BSM categories (Table 3), and the residual variance in short-term
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47 389 delivery was nearly identical between reduced and enlarged broods ($\sigma^2 = -0.001$, 95% CI =
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49 390 -0.002, 0.000). Overall, there was strong support that the effect of the BSM (difference
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4 391 between reduced and enlarged broods) on residual variance in short-term deliveries differed
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6 392 across years ($\sigma^2 = 0.22$, 95% CI = 0.01, 0.45).
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10 393 *Consequences of visit-level decisions for long-term provisioning*
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13 394 Long-term provisioning rates were similar across the two study years (Table 4, Figure 5a)
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15 395 and showed changes as a function of the BSM. In 2011, parents increased their average
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17 396 provisioning rates in response to the BSM (difference between enlarged versus reduced
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19 397 broods; $\beta = 0.83$, 95% CI = 0.22, 1.12), but there was only moderate support for an increase
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21 398 in provisioning rates across BSM in 2010 (difference between reduced and enlarged
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23 399 broods: $\beta = 0.31$, 95% CI = -0.10, 0.75) (Table 4).
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28 400 Long-term nest-level delivery increased significantly with increasing brood size ($\beta = 0.30$,
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30 401 95% CI = 0.16, 0.46). After controlling for brood size, there was moderate support for year-
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32 402 specific differences in long-term deliveries per nestling as a function of the BSM ($\beta = 0.46$,
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34 403 95% CI = -0.47, 1.63, Bayesian p-value = 0.14) (Table 4, Figure 5b). In 2010, there was
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36 404 moderate support for the interpretation that long-term delivery per nestling decreased
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38 405 across the BSM levels (difference between reduced and enlarged broods: $\beta = -0.52$, 95% CI
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40 406 = -1.29, 0.43, Bayesian p-value = 0.14). In contrast, there was no support for BSM-related
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42 407 differences in long-term delivery in 2011 (difference between reduced and enlarged broods:
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44 408 $\beta = 0.03$, 95% CI = -0.97, 1.09), suggesting that delivery per nestling did not differ as a
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46 409 function of the BSM in that year.
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5 411 **Discussion**
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8 412 Parents provisioning offspring can adopt a variety of (non-exclusive) tactics to cope with
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10 413 changes in brood demand. We used brood size manipulations in great tits and found that the
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12 414 tactics adopted by parents to cope with increased brood demand differed across two years.
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14 415 In 2010, there was no evidence of BSM-related shifts in mean IVIs or prey selectivity. As a
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16 416 result, there was moderate support that nestlings from enlarged broods experienced lower
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18 417 long-term delivery compared with nestlings from reduced broods. In contrast, in 2011,
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20 418 parents exhibited BSM-related shifts in both mean IVI and prey selectivity and there was
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22 419 no support for BSM-related differences in long-term delivery experienced by nestlings.
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24 420 Moreover, we found strong evidence that BSM-related differences in variance in IVI and in
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26 421 delivery per trip differed across years. Shifts in variances in IVIs and delivery per trip were
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28 422 consistent with stronger variance-sensitive behaviour in 2010 compared with 2011. The
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30 423 difference between years was not a variable we had considered going into the experiment,
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32 424 but the consequences are important and suggest a potential role of ecological conditions in
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34 425 mediating the tactics adopted by parents to cope with increased brood demand.
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41 426 *Are shifts in behaviour the result of shifts in ecology?*
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45 427 Parental responses to the brood size manipulations differed across the two study years. We
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47 428 interpret this as evidence that ecological conditions play an important role in shaping
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49 429 parental provisioning decisions. However, the same brood size manipulations in the two
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51 430 study years did not have the same effect on brood size on the day of filming (Table 1,
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53 431 Figure 3). In 2010, the net effect of the brood size manipulation was much smaller
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5 432 compared with 2011. Therefore, an alternative interpretation for the year-related differences
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7 433 in parental provisioning behaviour as a function of the BSM is that parents were responding
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9 434 to year-related differences in the effects of the BSM treatment. However, analyses using
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11 435 brood size on the day of filming, rather than BSM category, produced qualitatively similar
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14 436 results (see ESM Text S2 and Table S1), indicating that parents responded to variation in
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16 437 nestling demand (i.e. number of nestlings) differently across the two years, which supports
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18 438 our interpretation that parental provisioning decisions were shaped by year-specific
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21 439 ecological conditions.

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24 440 Although replication of ecological conditions is needed to draw strong conclusions
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26 441 regarding which specific aspects of the ecological conditions shaped the across year
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28 442 differences in parental provisioning decisions (Hurlbert, 1984), the comparison between the
29
30 443 two years is intriguing. Ecological conditions in the two study years may have played an
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32 444 important role in shaping the options available to parent great tits to cope with increasing
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34 445 brood demand. The 2010 season appears to have been a bad one for breeding great tits in
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36 446 our population (see Nicolaus et al., 2015, and ESM Text S3) and the more challenging
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38 447 ecological conditions may have limited the options available to parents. Average daytime
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40 448 temperatures in our study area were colder compared with 2011 (ESM Figure S2a), and
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42 449 rainfall was more frequent (ESM Figure S2b). There was also higher nestling mortality in
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44 450 2010 (ESM Figure S3), suggesting that parents struggled to meet the energy requirements
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46 451 of their broods in that year.

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48 452 In the year with these more challenging conditions (2010), caterpillars, the preferred prey
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53 453 of great tits (Naef-Daenzer et al., 2000; Wilkin et al., 2009), made up a substantially smaller
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4 454 proportion of the diet compared with the good year (2011) (Table 2). This could reflect a
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6 455 lower availability of caterpillars in 2010, or may indicate that parents were overall less
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8 456 selective to compensate for higher brood requirements in that year, for example if the lower
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10 457 temperature increased the energy requirements of nestlings via increased thermoregulatory
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12 458 costs (Sanz and Tinbergen, 1999). The latter seems unlikely, because if that were the case,
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14 459 we would have expected a higher overall variance in prey type or size in 2010 compared
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16 460 with 2011, which would have been expected from reduced selectivity and a reduction in the
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18 461 minimum quality/size threshold for accepting prey items. We would expect that parent great
19
20 462 tits would still accept large and/or high quality prey when encountered, resulting in a wider
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22 463 range of prey sizes being accepted, thereby reducing the mean but increasing the variance
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24 464 in size of provisioned items. In contrast, we observed that variance in load size was lower
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26 465 in 2010 compared with 2011 (see Table 3, Figure 4e). Therefore, differences in the size and
27
28 466 type of prey provisioned across the two study years probably reflect a lower availability of
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30 467 large, preferred prey types in 2010.

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33 468 There was only moderate support for an increase in long-term average provisioning rates as
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35 469 a function of the BSM in 2010. This increase in long-term average provisioning rate was
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37 470 not accompanied by any BSM-related differences in either mean IVI or in prey type (Table
38
39 471 2). There was moderate support for an increase in load size with increasing brood demand.
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41 472 Taken together, the tendency to increase long-term provisioning rates and mean load size
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43 473 per visit across the BSM categories suggests that parent great tits may have attempted to
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45 474 meet increased brood demand in 2010 via increased provisioning effort.
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4 475 In 2011, there was strong support for BSM-related difference in mean IVI, but no support
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7 476 for BSM-related differences in load size. Additionally, there was strong support for prey
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9 477 selectivity being greatest for parents provisioning reduced broods (i.e. they provisioned a
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11 478 significantly greater fraction of preferred prey to the young). Thus, the BSM-related
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14 479 increase in long-term average provisioning rates in 2011 came about via a combination of
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16 480 increased effort (BSM-related changes in IVI) and a decrease in selectivity (for enlarged
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18 481 broods). The greater overall increase in provisioning rates in 2011 compared with 2010
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21 482 may have occurred because the more favourable ecological conditions associated with that
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23 483 year meant that parents had more options available for increasing provisioning rates. The
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25 484 greater overall availability of caterpillars in 2011 implied that increased prey selectivity
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28 485 was a viable option for parents provisioning reduced broods in that year, and the more
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30 486 favourable temperatures may have meant that parents did not require as much time for self-
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33 487 feeding or brooding the young for thermoregulation, making it possible to increase long-
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35 488 term provisioning effort for enlarged broods to a greater extent compared with parents
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37 489 provisioning enlarged broods in the more challenging year (2010).

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41 490 There was moderate support for long-term average delivery experienced by nestlings
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43 491 varying as a function of the BSM in 2010, but there was no support for such an effect in
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45 492 2011. In 2010, the more challenging year in terms of ecological conditions, delivery to the
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48 493 nest decreased as a function of experimentally increased brood demand after controlling for
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50 494 current brood size. This indicates that the per nestling delivery was lower in enlarged
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52 495 broods, and suggests that the brood size manipulation did indeed place nestlings from
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55 496 enlarged broods on a negative energy budget relative to nestlings from control and reduced

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4 497 broods. This was not the case in 2011, when parents apparently fully compensated for
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6 498 increased brood demand as there was no evidence for BSM-related differences in long-term
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9 499 average delivery after controlling for number of nestlings. Analyses of nestling mortality
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11 500 rates as a function of the BSM in each year corroborate our interpretation that the BSM
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13 501 produced meaningful changes in the probability of energy shortfall in 2010 but not in 2011.
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16 502 In 2010, nestling mortality increased significantly across BSM categories, but there were no
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19 503 BSM-related differences in nestling mortality in 2011 (ESM Table S2).

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22 504 ***Mechanisms underlying variance in provisioning behaviour***

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25 505 We observed year-specific BSM-related differences in residual variance in parental inter-
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27 506 visit intervals (IVIs) and in short-term delivery rates experienced by nestlings. The results
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30 507 were consistent with 2010 exerting stronger variance-sensitive provisioning compared with
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32 508 2011. In 2010, residual variance in parental IVIs tended to be smallest for reduced broods
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34 509 and greatest for enlarged broods. More importantly, the combined effects of both parents
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37 510 resulted in significantly greater residual variance in short-term delivery experienced per
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39 511 nestling in enlarged compared with reduced broods. Our study was not designed to assess
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41 512 the potential mechanism(s) that parent great tits used that may have caused differences in
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43 513 the variance in their prey encounter rates. Nevertheless, the results here suggest that this did
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46 514 not come about via changes in prey selectivity (either size or type, Table 2). In contrast to
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49 515 our results, two previous experimental studies of variance-sensitive provisioning found that
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51 516 brood demand-related shifts in the variance in prey delivery were associated with changes
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53 517 in the size and type of prey delivered (Moore, 2002; Whittingham and Robertson, 1993).
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56 518 Differences in the foraging ecologies of the species involved may explain the different
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4 519 findings. Both previous reports of brood-size dependent variance-sensitive provisioning
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6 520 involved species where parents could choose between foraging in alternative, discrete
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8 521 habitat types that differed in prevailing prey type (e.g. common terns foraging in a small
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10 522 pond or large lake: Moore, 2002; red-winged blackbirds foraging in woodland or marsh:
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12 523 Whittingham and Robertson, 1993; see Ydenberg, 2008). Great tits, however, forage in
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14 524 what appears to be a single habitat type, deciduous forest. While it is likely that prey are
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16 525 distributed unevenly across this habitat and that parent great tits are making use of this
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18 526 heterogeneity in their decision-making (Naef-Daenzer, 2000), we cannot assess this with
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20 527 our data.
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26 528 If great tits do not exhibit shifts in prey selectivity or foraging habitat, how were BSM-
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28 529 related differences in the variance in prey delivery rates achieved? One possibility is that
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30 530 great tits strategically adjust their patch departure decisions (Stephens and Charnov, 1982).
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32 531 However, this has only been explicitly considered for cases where animals obtain
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34 532 diminishing returns in a patch (i.e. consume/load multiple prey items per patch), and it is
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36 533 unclear that this can be extended to single prey loaders such as great tits (and with no
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38 534 obvious evidence of prior patch depletion from parental self-feeding). Detailed
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40 535 observations of parents on provisioning trips (e.g. when using feeders) may provide some
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42 536 insights into how BSM-related differences in variance are achieved.
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48 537 Interestingly, BSM-related differences in the variance in IVIs in 2011 ran counter to the
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50 538 patterns observed in 2010: variance in provisioning behaviour decreased with increasing
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52 539 brood demand. We suggest that the patterns of decreasing variance in parental IVIs
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54 540 observed in 2011 is not the result of variance-sensitive provisioning decisions by parents
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4 541 for two reasons. First, the observed shifts in mean parental behaviour in 2011 meant that
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6 542 there were no treatment-related differences in average long-term delivery experienced per
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9 543 nestling. Since all nestlings in all treatments appear to have received the same long-term
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11 544 rate of food delivery, and since there were no BSM-related differences in mortality,
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13 545 nestlings from all BSM treatments should have been at a similar location on the utility
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15 546 curve in Figure 2, and hence no BSM-related changes in preference (or aversion) for
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17 547 variance were expected.
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22 548 Second, residual variance in short-term delivery experienced at the level of nestlings did
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24 549 not vary as a function of the BSM in 2011. Variance-sensitive behaviour by parents is
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26 550 expected when deviations in energy above and below the mean energy delivery have
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28 551 asymmetric fitness consequences for offspring (Figure 2). Variance-sensitive responses will
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30 552 only have fitness consequences for offspring if the combined effect of each parent's
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32 553 individual response to variation in brood demand results in shifts in the variance in energy
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34 554 delivery experienced at the level of the nestling. Thus, the lack of BSM-related differences
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36 555 in variance in nestling-level delivery suggests that the BSM-related patterns in parental
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38 556 IVIs in that year were not due to parents adaptively adjusting the variance in delivery in
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40 557 response to non-linear utility functions of nestlings. Nonetheless, we cannot rule out that
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42 558 parents did exhibit adaptive variance-sensitive responses to variation in brood demand, but
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44 559 that these effects were slight in 2011 compared with the ones observed in 2010 and
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46 560 therefore masked by other processes influencing variance in parental IVIs.
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53 561 What processes might have generated the observed pattern of decreasing variance in IVIs
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55 562 with increasing brood demand in 2011? One possible explanation is that the decrease in
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4 563 variance with brood size may have been a by-product of parental strategic decisions to
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6 564 decrease prey selectivity as a way of increasing mean energy delivery (via reduced IVIs)
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9 565 (see Weissburg, 1991 for worked examples). However, this seems unlikely as additional
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11 566 analyses of our data (not shown) demonstrated that residual variances in IVI did not differ
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14 567 as a function of prey type in the present study. Alternatively, parents may divide their time
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16 568 between parenting and other activities in a more variable manner between trips when brood
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18 569 demand is reduced (see also Westneat et al., 2017). For example, parents with small broods
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21 570 in good years may be able to forage intensively for a while close to the nest, and once their
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23 571 brood has been adequately fed they are free to devote some time to territory defense
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25
26 572 (Markman et al., 1995) and/or foraging for themselves (Martins and Wright, 1993). This
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28 573 could be a sensible strategy if territory defense and/or sampling new and different foraging
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30 574 patches requires additional travel further from the nest and is therefore best done separately
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32
33 575 from any continuous provisioning back at the nest. In this way, provisioning rates would
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35 576 decrease and simultaneously the variance in IVI would be higher in reduced broods, but
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37 577 only in good years. However, a more comprehensive quantification of parental behaviour
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39 578 (e.g. detailed observations on the time allocated to provisioning versus other activities) as a
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41
42 579 function of year and BSM are obviously required to establish how changes in parental
43
44 580 effort affect mean provisioning rates and variance in IVIs. Additionally, field experiments
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46 581 that manipulate variability in food reward (e.g. Ratikainen et al., 2010) would be useful to
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49 582 disentangle the effects of multiple co-occurring processes on residual variance in parental
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52 583 provisioning.

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4 584 *Conclusions*
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8 585 We found that great tit parents respond in several ways to changes in brood demand, and
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10 586 that their responses likely depend on a complex mix of factors. Differing ecological
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12 587 conditions are the most likely reason for the year-specific parental responses. Parents faced
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14 588 with increased brood sizes increased long-term average visit rates in both years, but the
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16 589 increase was greater in the good year (2011) than in the bad (2010). In 2010, the increased
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18 590 long-term average provisioning rate was consistent with an increase in provisioning effort,
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20 591 while in 2011 it was consistent with a combination of increased effort and reduced
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22 592 selectivity. Great tit parents exhibited behaviour consistent with predictions from variance-
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24 593 sensitive foraging theory, but only in the bad year (2010). This supports an earlier claim
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26 594 that variance-sensitive decisions are favoured most strongly under more challenging
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28 595 ecological conditions (Moore, 2002). One of the most exciting implications of having
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30 596 manipulated brood size over two years with such different ecological conditions is that it
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32 597 revealed the intriguing possibility that conditions can strongly affect parental tactics for
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34 598 coping with increased brood demand. For example, decisions regarding the use of one
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36 599 strategy (e.g. shifting energetic quantity of prey at the expense of their nutritional quality)
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38 600 might influence the value of then also adopting other strategies (e.g. working harder or
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40 601 variance-sensitivity) (see also Ratikainen et al., 2012). The potential for such overlapping
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42 602 and interactive effects of the different strategic options involving both the means and
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44 603 variances in rewards highlights the need to adopt a more systems-level approach to the
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46 604 study of parental care decisions (i.e. studying integrated, multi-trait phenotypes) across an
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48 605 array of ecological conditions.
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8 607 **Data accessibility**

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11 608 All data used in the analyses presented in this paper will be deposited on Dryad upon
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13 609 publication of the paper.
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20 611 **References**

21
22
23 612 Bates D, Maechler M, Bolker B, Walker S, 2015. Fitting linear mixed-effects models using
24
25 613 lme4. *Journal of Statistical Software* 67:1-48. doi: 10.18637/jss.v067.i01.

26
27
28 614 Bateson M, Kacelnik A, 1996. Rate currencies and the foraging starling: The fallacy of the
29
30 615 averages revisited. *Behav Ecol* 7:341-352. doi: 10.1093/beheco/7.3.341.

31
32
33 616 Cleasby I, Nakagawa S, 2011. Neglected biological patterns in the residuals. *Behav Ecol*
34
35 617 *Sociobiol* 65:2361-2372. doi: 10.1007/s00265-011-1254-7.

36
37 618 Clutton-Brock TH, 1991. *The evolution of parental care*. Princeton, New Jersey: Princeton
38
39 619 University Press.

40
41
42 620 Cohen J, 1990. Things I have learned (so far). *American Psychologist* 45:1304-1312.

43
44 621 Cumming G, Finch S, 2005. Inference by eye: confidence intervals and how to read
45
46 622 pictures of data. *American Psychologist* 60:170-180.

47
48
49 623 Gelman A, Su Y-S, 2015. arm: Data analysis using regression and multilevel/hierarchical
50
51 624 models. R package version 1.8-6.

52
53
54 625 Hadfield J, 2010. MCMC Methods for multi-response generalized linear mixed models: the
55
56 626 MCMCglmm R Package. *Journal of Statistical Software* 33:1-22.
57
58
59
60

- 1
2
3
4 627 Hurlbert SH, 1984. Pseudoreplication and the design of ecological field experiments. *Ecol*
5
6 628 *Monogr* 54:187-211.
7
8
9 629 Leonard ML, Horn AG, 1998. Need and nestmates affect begging in tree swallows. *Behav*
10
11 630 *Ecol Sociobiol* 42:431-436. doi: 10.1007/s002650050457.
12
13
14 631 Lifjeld JT, 1989. Central place foraging: optimal load size for net and gross energy
15
16 632 maximizers. *Oikos* 55:397-401. doi: 10.2307/3565600.
17
18
19 633 Markman S, Yomtov Y, Wright J, 1995. Male parental care in the orange-tufted sunbird:
20
21 634 behavioral adjustments in provisioning and nest guarding effort. *Anim Behav*
22
23 635 50:655-669. doi: 10.1016/0003-3472(95)80127-8.
24
25
26 636 Martins TLF, Wright J, 1993. Cost of reproduction and allocation of food between parent
27
28 637 and young in the swift (*Apus apus*). *Behav Ecol* 4:213-223. doi:
29
30 638 10.1093/beheco/4.3.213.
31
32
33 639 Moore DJ, 2002. The provisioning tactics of parent common terns (*Sterna hirundo*) in
34
35 640 relation to brood energy requirement. Burnaby: Simon Fraser University, p. 418.
36
37
38 641 Moreno J, Cowie RJ, Sanz JJ, Williams RSR, 1995. Differential response by males and
39
40 642 females to brood manipulations in the pied flycatcher: energy expenditure and
41
42 643 nestling diet. *J Anim Ecol* 64:721-732.
43
44
45 644 Mutzel A, Olsen A-L, Araya-Ajoy YG, Mathot KJ, Nicolaus M, Wijmenga JJ, Wright J,
46
47 645 Kempenaers B, Dingemanse NJ, *In preparation*. Personality-related differences in
48
49 646 offspring provisioning under threat of nest predation. *Behav Ecol*.
50
51
52 647 Naef-Daenzer B, 2000. Patch time allocation and patch sampling by foraging great and blue
53
54 648 tits. *Anim Behav* 59:989-999.
55
56
57
58
59
60

- 1
2
3
4 649 Naef-Daenzer L, Naef-Daenzer B, Nager RG, 2000. Prey selection and foraging
5
6 performance of breeding Great Tits *Parus major* in relation to food availability. J
7 650
8 Avian Biol 31:206-214.
9 651
10
11 652 Nakagawa S, Cuthill IC, 2007. Effect size, confidence interval and statistical significance: a
12
13 practical guide for biologists. Biol Rev 82:591-605. doi: 10.1111/j.1469-
14 653
15 185X.2007.00027.x.
16 654
17
18 655 Nakagawa S, Schielzeth H, 2017. Coefficient of determination R² and intra-class
19
20 correlation coefficient ICC from generalized linear mixed-effects models revisited
21 656
22 and expanded. bioRxiv. doi: 10.1101/095851.
23 657
24
25 658 Nicolaus M, Mathot KJ, Araya-Ajoy YG, Mutzel A, Wijmenga JJ, Kempenaers B,
26
27 Dingemans NJ, 2015. Does coping style predict optimization? An experimental test
28 659
29 in a wild passerine bird. Proceedings of the Royal Society B Biological Sciences
30 660
31 282:20142405. doi: 10.1098/rspb.2014.2405.
32 661
33
34 662 Author. 2015. R: A language and environment for statistical computing. R Foundation for
35
36 Statistical Computing. Vienna, Austria: See <http://www.R-project.org>.
37 663
38
39 664 Ratikainen II, Sødal LRH, Kazem AJN, Wright J, 2012. Private and public information use
40
41 strategies by foraging groups of wild Siberian jays. Anim Behav 83:1243-1251.
42 665
43
44 666 Ratikainen II, Wright J, Kazem AJN, 2010. Social class influences degree of variance
45
46 sensitivity in wild Siberian jays. Behav Ecol 21:1067-1072.
47 667
48
49 668 Royle NJ, Smiseth PT, Kölliker M, 2012. The evolution of parental care. Oxford: Oxford
50
51 University Press.
52 669
53
54
55
56
57
58
59
60

- 1
2
3
4 670 Sanz JJ, Tinbergen JM, 1999. Energy expenditure, nestling age, and brood size: an
5
6
7 671 experimental study of parental behavior in the great tit *Parus major*. *Behav Ecol*
8
9 672 10:598-606. doi: 10.1093/beheco/10.5.598.
10
11 673 Stephens DW, Charnov E, 1982. Optimal foraging: Some simple stochastic models. *Behav*
12
13 674 *Ecol Sociobiol* 10:251-263. doi: 10.1007/BF00302814.
14
15
16 675 Sutherland WJ, Anderson CW, 1987. Six ways in which a foraging predator may encounter
17
18 676 options with different variances. *Biol J Linn Soc* 30:99-114. doi: 10.1111/j.1095-
19
20 677 8312.1987.tb00292.x.
21
22
23 678 Templeton AR, Lawlor LR, 1981. The Fallacy of the Averages in Ecological Optimization
24
25 679 Theory. *The American Naturalist* 117:390-393.
26
27
28 680 Wasserstein RL, Lazar NA, 2016. The ASA's statement on p-values: context, process, and
29
30 681 purpose. *The American Statistician* 70:129-133. doi:
31
32 682 10.1080/00031305.2016.1154108.
33
34
35 683 Weissburg M, 1991. Mean-variance sets for dietary choice models: simplicity in a complex
36
37 684 world. *Evolutionary Ecology* 5:1-11.
38
39
40 685 Welsh AH, Peterson AT, Altmann SA, 1988. The fallacy of averages. *Am Nat* 132:277-288.
41
42 686 doi: 10.1086/284850.
43
44
45 687 Westneat DF, Hatch MI, Wetzel DP, Ensminger AL, 2011. Individual variation in parental
46
47 688 care reaction norms: integration of personality and plasticity. *Am Nat* 178:652-667.
48
49
50 689 Westneat DF, Mutzel A, Bonner S, Wright J, 2017. Experimental manipulations of brood
51
52 690 size affects several levels of phenotypic variance in offspring and parent pied
53
54 691 flycatchers. *Behav Ecol Sociobiol*:In press.
55
56
57
58
59
60

- 1
2
3
4 692 Westneat DF, Schofield M, Wright J, 2013. Parental behavior exhibits among-individual
5
6 693 variance, plasticity, and heterogeneous residual variance. *Behav Ecol* 24:598-604.
7
8 694 doi: 10.1093/beheco/ars207.
9
10
11 695 Westneat DF, Wright J, Dingemanse NJ, 2015. The biology hidden inside residual within-
12
13 696 individual variance. *Biol Rev (Camb)* 90:729-743.
14
15
16 697 Whittingham LA, Robertson RJ, 1993. Nestling hunger and parental care in red-winged
17
18 698 blackbirds. *Auk* 110:240-246.
19
20
21 699 Wilkin TA, King LE, Sheldon BC, 2009. Habitat quality, nestling diet, and provisioning
22
23 700 behaviour in great tits *Parus major*. *J Avian Biol* 40:135-145.
24
25
26 701 Williams GC, 1966. Natural selection, the costs of reproduction, and a refinement of Lack's
27
28 702 principle. *The American Naturalist* 100:687-690. doi: 10.2307/2459305.
29
30
31 703 Winkler DW, 1987. A general model for parental care. *The American Naturalist* 130:526-
32
33 704 543. doi: 10.2307/2461702.
34
35 705 Wright J, 1998. Helpers-at-the-nest have the same provisioning rule as parents:
36
37 706 experimental evidence from play-backs of chick begging. *Behav Ecol Sociobiol*
38
39 707 42:423-429. doi: 10.1007/s002650050456.
40
41
42 708 Wright J, Both C, Cotton PA, Bryant D, 1998. Quality vs. quantity: energetic and
43
44 709 nutritional trade-offs in parental provisioning strategies. *J Anim Ecol* 67:620-634.
45
46
47 710 Wright J, Leonard ML, 2010. The evolution of begging: competition, cooperation &
48
49 711 communication. Dordrecht: Kluwer Academic Publishers.
50
51
52 712 Ydenberg RC, 1994. The behavioral ecology of provisioning in birds. *Ecoscience* 1:1-14.
53
54 713 Ydenberg RC, 2008. Provisioning. In: Stephens DW, Brown JL, Ydenberg RC, editors.
55
56 714 Foraging Chicago: University of Chicago Press. p. 273-303.
57
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715 **Figure legends:**

716 **Figure 1:** Graphical illustration of the predicted outcomes from three non-mutually
717 exclusive strategies for coping with changes in brood demand (R= reduced, C = control,
718 and E = enlarged broods). Predicted effects of brood demand on means are illustrated with
719 solid black lines (directional prediction) or grey filled wedges (when the predicted response
720 can vary from no change to a directional effect). Solid grey lines reflect cases where there is
721 no *a priori* predicted directional effect on means. Predicted effects on variances are
722 illustrated with dotted lines where *a priori* predictions exist. All else being equal, a) when
723 parents increase effort, increased energy delivery to the nest arises via decreased IVIs
724 and/or increased energy per visit. This may be achieved via increased prey selectivity, or by
725 increasing the time/energy devoted to provisioning. b) When parents trade off energetic
726 quantity versus nutritional quality, increased energy delivery to the nest arises via decreased
727 IVIs and/or increased energy delivery per visit, which are the result of decreasing prey
728 selectivity (or prey shifts). c) When parents display variance sensitivity, variance in the rate
729 of energy delivery is predicted to increase with increasing brood demand via an increase in
730 variance in IVI and/or an increase in variance in prey size. There is no universal prediction
731 regarding the effect of brood demand on mean IVI or mean energy delivered per visit when
732 animals exhibit variance-sensitivity. Similarly, there is no universal prediction for how
733 variance sensitivity should affect prey selectivity (in some cases, higher variance arises
734 through increased selectivity, in other cases it can be achieved through decreased
735 selectivity). Effects on means depend on characteristics of the prey (e.g. energy content,
736 encounter rate, etc.).

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5 737 **Figure 2:** Illustration of the non-linear relationship between food delivery and offspring
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7 738 fitness, which forms the basis of the theoretical argument for variance sensitivity of parents
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9 739 to prey encounter rates when feeding offspring. On the left side of the graph, a positive
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11 740 deviation (black whisker) from the mean food delivery (blue circle) has a larger benefit
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13 741 compared with the cost of an equal magnitude negative deviation ($b > c$) (e.g. when
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15 742 nestlings are in poor condition). Thus, parents are expected to be variance-prone. The right
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17 743 side of the graph illustrates a scenario where a positive deviation above the mean delivery
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19 744 (red circle) has a smaller benefit compared to the cost of an equal magnitude negative
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21 745 deviation ($c > b$) (e.g. when nestlings are in good condition), hence, parents should be
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23 746 variance-averse. Note that we assume that brood size manipulations will result in
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25 747 systematic reductions in the amount of food per nestling per unit of time, leading to an
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27 748 increased likelihood that parents become variance-prone.

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33 749 **Figure 3:** Brood size on day of filming for reduced (R), control (C) and enlarged (E)
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35 750 broods. Values presented are means \pm 1 s.e. calculated from raw data.

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39 751 **Figure 4:** Effects of brood size manipulation on (I) means (top row) and (II) variances
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41 752 (bottom row) in visit-level provisioning behavior. Shown are: a) and d) inter-visit intervals
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43 753 (IVIs), b) and e) load size, c) & f) short-term delivery (load size/inter-feed interval/brood
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45 754 size). Note that statistical analyses of data in c) & f) were of total short-term delivery with
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47 755 current brood size as a covariate. Top panel depicts means \pm 1 s.e. (calculated from raw
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49 756 data). The lower panel illustrates the residual variances estimates ($\sigma \pm$ 95% CI) extracted
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51 757 from models (see methods for further details). Data from 2010 is shown with filled circles,
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53 758 and data from 2011 is shown with hollow circles.

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4 759 **Figure 5:** Effects of brood size manipulations (BSMs) on measures of long-term
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6 760 provisioning. Shown are: a) long-term provisioning rate (number of visits by each parent
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8 761 per 30 minute observation) and b) long-term nestling level delivery (sum of load sizes
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10 762 delivered by both parents per 30 minute observation divided by brood size. Note that
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12 763 statistical analyses of long-term nestling level delivery included current brood size as a
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14 764 covariate, but ratios (long-term delivery/brood size) are used here for illustrative purposes.
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16 765 In b) delivery/nestling is used for illustrative purposes. Values shown are means \pm 1 s.e.
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19 766 calculated from raw data.
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3 767 **Table 1:** Brood size (square-root (“Sqrt”) transformed) on the day of the provisioning
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5 768 observations (nestling age 12) as a function of initial clutch size and the brood size manipulation
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7 769 (BSM) for each year (see Fig. 3 for non-transformed brood sizes). Estimated means (β) or
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9 770 variances (σ^2) are reported with their 95% credible intervals (CI).
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Sqrt brood size	
<i>Fixed effects</i>	β (95% CI)
2010	
Reduced	1.94 (1.75, 2.18)
Control	2.29 (2.10, 2.54)
Enlarged	2.63 (2.42, 2.84)
2011	
Reduced	2.15 (1.94, 2.34)
Control	2.63 (2.42, 2.87)
Enlarged	3.17 (2.96, 3.36)
Natural clutch size ¹	0.06 (0.01, 0.11)
<i>Random effects</i>	$\sigma^2 \pm 95\% \text{ CI}$
Plot	0.06 (0.01, 0.14)
Residual	0.07 (0.05, 0.10)

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772 ¹Clutch size centered within year.

773 **Table 2:** Effects of brood size manipulations (BSM) on means of visit-level provisioning behaviour. Variation in parental decisions
 774 [mean inter-visit interval (IVI, in seconds), load size (bill volumes), and prey type (preferred or non-preferred)], the net consequence
 775 of combined parental decisions for nestlings [short-term delivery (prey size/IFI, controlling for brood size)] in relation to the brood
 776 size manipulation in each year. IVI and delivery data were ln transformed, and load size was square root transformed. Data were
 777 modelled with Gaussian (IVI, prey size, delivery) or binomial (prey type) errors. Values presented are estimated means (β) and their
 778 95% credible intervals (CI).

Fixed effects	Parental visit-level behavior			Ln(Nestling-level Delivery)
	Ln IVI	Sqrt load size	Prey type ¹	
	$\beta \pm 95\% \text{ CI}$	$\beta \pm 95\% \text{ CI}$	$\beta \pm 95\% \text{ CI}$	$\beta \pm 95\% \text{ CI}$
Brood size	NA	NA	NA	0.12 (0.07, 0.18)
2010				
Reduced	4.90 (4.70, 5.09)	0.92 (0.86, 0.98)	-1.83 (-2.39, -1.25)	-5.24 (-5.52, -4.96)
Control	4.85 (4.64, 5.05)	0.97 (0.91, 1.03)	-1.96 (-2.55, -1.30)	-5.14 (-5.56, -4.80)
Enlarged	4.85 (4.66, 5.06)	1.00 (0.94, 1.06)	-2.05 (-2.67, -1.45)	-5.06 (-5.47, -4.62)
2011				
Reduced	5.19 (4.95, 5.40)	1.16 (1.10, 1.23)	1.47 (1.02, 2.22)	-5.13 (-5.47, -4.76)
Control	5.09 (4.87, 5.30)	1.14 (1.07, 1.22)	0.61 (0.06, 1.21)	-5.18 (-5.64, -4.78)

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3 Enlarged 4.88 (4.70, 5.06) 1.18 (1.11, 1.23) 0.74 (0.20, 1.24) -5.11 (-5.70, -4.53)
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9 780 Footnotes

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11 781 1. Prey type was modelled with binomial errors. Preferred prey (i.e. caterpillars) were coded as 1, and non-preferred prey (i.e. others)
12 were coded as 0. Estimates are in ln (odds ratios): therefore, an estimate of 0 represents an equal proportion of preferred and non-
13 preferred prey in the diet; an estimate of 0.69 represents cases where preferred prey are twice as frequent as non-preferred prey, and an
14 estimate of -0.69 represents cases where non-preferred prey are provisioned twice as much as preferred prey.
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3 **Table 3:** Effects of brood size manipulations (BSM) on variances of visit-level provisioning
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6 786 behaviour, obtained from the same models as in Table 2. Variance components of inter-visit
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8 787 interval (IVI, in seconds), load size (bill volumes), and short-term delivery (prey size/IFI,
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10 788 controlling for brood size) in relation to the brood size manipulation in each year. IVI and
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12 789 delivery data were ln transformed and load size was square-root transformed. Values presented
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15 790 are residual variances (σ^2) and their 95% credible intervals (CI).

Treatment	Parent-level behaviour		Ln (Nestling-level
	Ln IVI	Sqrt load size	Delivery)
	$\sigma^2 \pm 95\% \text{ CI}$	$\sigma^2 \pm 95\% \text{ CI}$	$\sigma^2 \pm 95\% \text{ CI}$
2010			
Reduced	0.55 (0.46, 0.64)	0.05 (0.04, 0.06)	0.88 (0.73, 1.04)
Control	0.56 (0.47, 0.67)	0.06 (0.05, 0.07)	1.04 (0.89, 1.22)
Enlarged	0.60 (0.51, 0.69)	0.05 (0.04, 0.05)	1.11 (0.96, 1.28)
2011			
Reduced	0.77 (0.61, 0.92)	0.09 (0.07, 0.10)	1.06 (0.87, 1.29)
Control	0.63 (0.52, 0.75)	0.09 (0.07, 0.10)	1.15 (0.95, 1.37)
Enlarged	0.57 (0.50, 0.65)	0.08 (0.07, 0.09)	1.04 (0.92, 1.16)

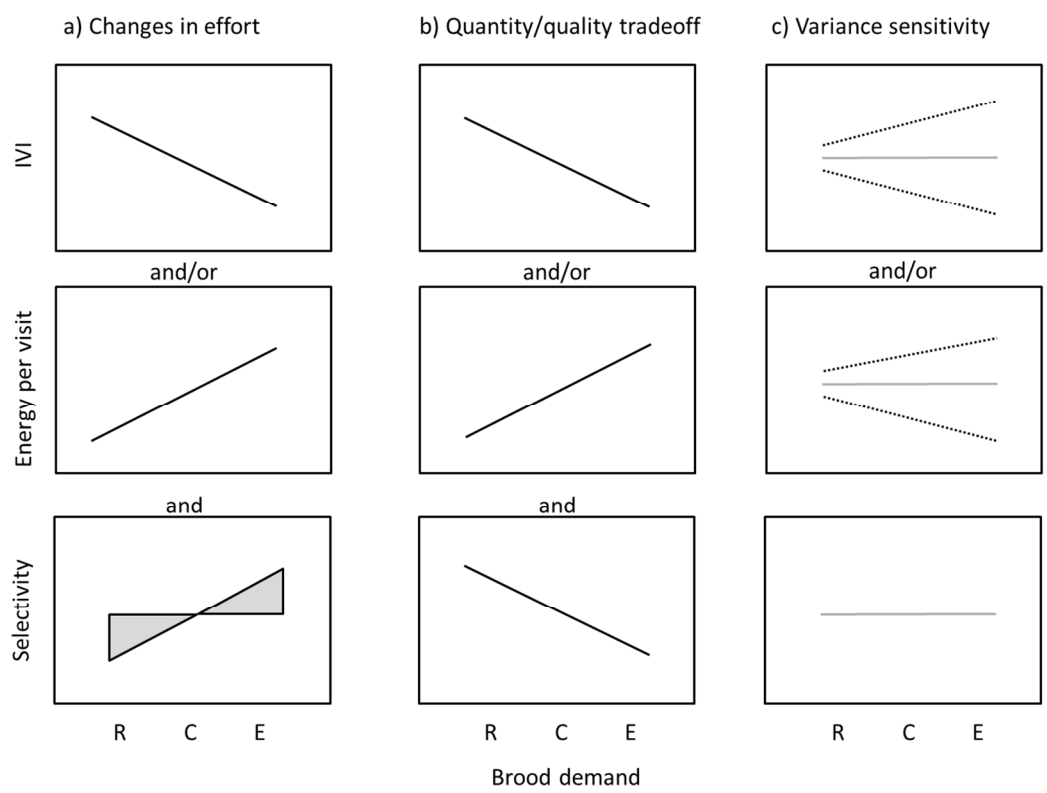
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4 792 **Table 4:** Long-term consequences of visit-level decisions: long-term parental provisioning rate
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6 793 (ln of visits per 30 minute observation), and long-term delivery per nestling (sum of load sizes
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8 794 per 30 minute observation, controlling for brood size).
9

	Long-term provisioning rate	Long-term delivery
	per parent	per nestling
<i>Fixed effects</i>	$\beta \pm 95\% \text{ CI}$	$\beta \pm 95\% \text{ CI}$
Brood size	NA	0.30 (0.16, 0.46)
2010		
Reduced	1.55 (1.21, 1.80)	3.36 (2.94, 3.99)
Control	1.67 (1.36, 2.03)	3.29 (2.75, 3.82)
Enlarged	1.86 (1.48, 2.13)	3.01 (2.53, 3.61)
2011		
Reduced	1.52 (1.13, 1.76)	4.27 (3.53, 4.75)
Control	1.65 (1.26, 1.92)	3.65 (3.21, 4.29)
Enlarged	2.12 (1.90, 2.50)	4.19 (3.58, 4.81)

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796 **Figure 1**

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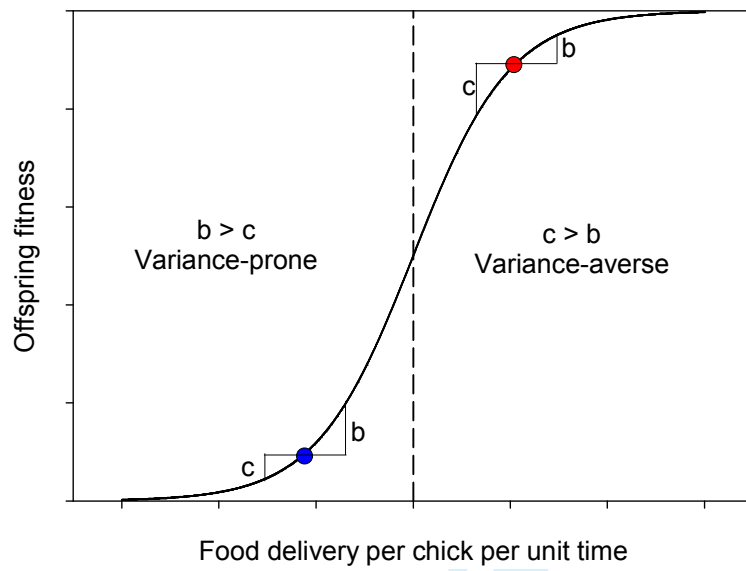
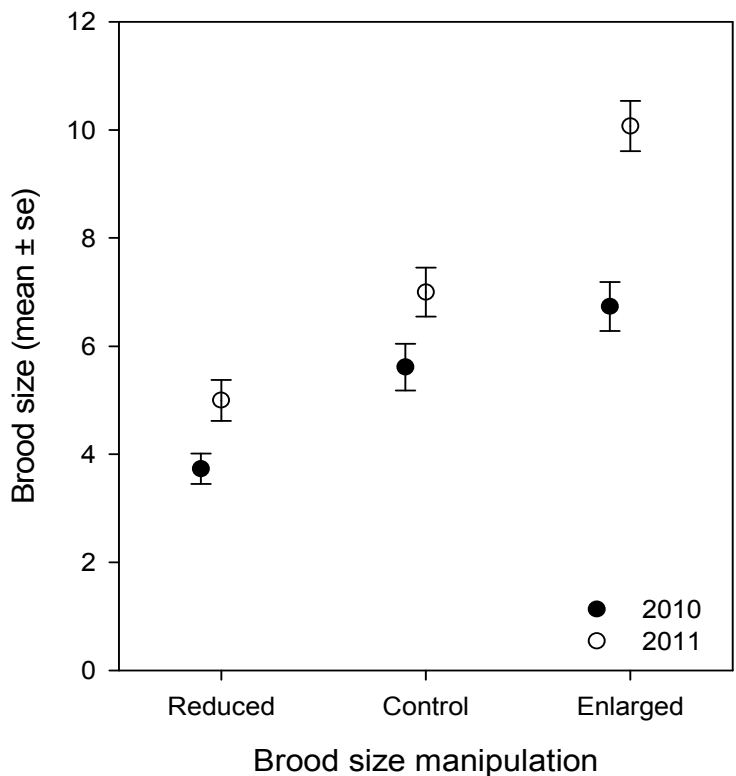


Figure 2

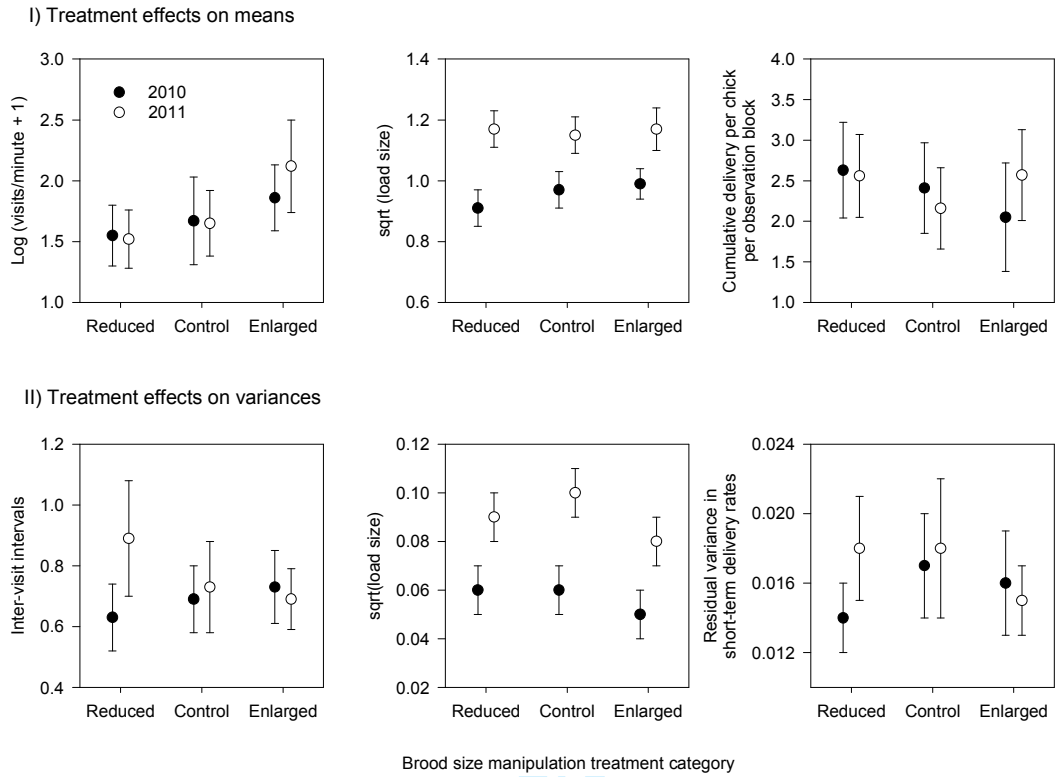
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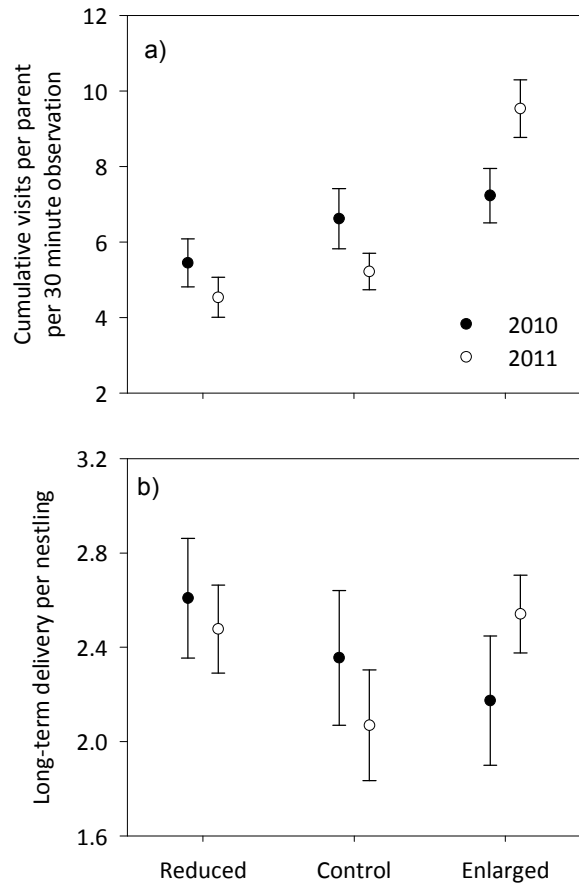
800 **Figure 3**

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802 **Figure 4**



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804 **Figure 5**

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19 **Text S1: Parameter estimation method**

20 We ran each analysis for a range of prior settings to ensure that our estimates were not
21 influenced by the choice of prior. The different prior settings produced very similar estimates for
22 all the parameters estimated (results not shown); analyses presented in the main text use an
23 inverse wishart prior where $\nu = 0.002$. Models were run for 103,000 iterations, with a burn-in
24 period of 3000 and thinning interval of 100. This produced a sample of 1000 estimates for each
25 model. These estimates were used to calculate the most likely value for each parameter (the
26 mode of the distribution), as well as its 95% credible interval.

27

28 **Text S2: Year differences in parental responses to current brood size**

29 The brood size manipulation (BSM) had different effects on brood size on the day of filming in
30 the two study years (main text Table 1 and Figure 3). Therefore, the year-related differences in
31 parental response to BSM could be due to either to year-specific ecology condition, or year-
32 specific effects of the BSM on realized brood size. To address these two possibilities, we
33 repeated the analyses presented in Tables 2 and 3 of the main text using as predictor variables
34 brood size on the day of filming (instead of BSM), year, and the brood size: year interaction. If
35 among-year differences in response to the BSM were due to the fact that the BSM had different
36 effects on brood size in the two years, then we expect brood size to predict parental behaviour,
37 but we do not expect a year x brood size interaction effect. Because average brood size differed
38 across the two study years, we centered brood size within-years prior to analyses to break the
39 correlation between year and brood size.

40 In all cases where we interpreted support for year-specific effects of the BSM on mean levels of
41 parental provisioning (IVI, load size, prey type) in the main text, analyses using current brood
42 size, year, and their interaction (rather than year-specific BSM) showed moderate to significant
43 support for the year x brood size interaction (Bayesian p-values: IVI = 0.062, load size = 0.041,
44 prey type = 0.017) (see Table S1). Further, consistent with analyses using a variable combining
45 year and BSM as a fixed effect (YBSM, a fixed effect with 6 levels, see Table 3, main text),
46 there was no support for an interaction between year and brood size on the nestling level delivery
47 (Bayesian p-value = 0.27). Modelling brood size as a fixed effect rather than the Year: BSM also
48 had no effect on the estimated residual variances relative to the analyses presented in the main
49 text. These findings indicate that between-year differences in the realized effect of the BSM on
50 the number of nestlings cannot account for year-related differences in the effect of the BSM
51 categories on parental provisioning decisions, and corroborate the interpretation that year-
52 specific ecological conditions shape the response of parent great tits to variation in brood
53 demand.

54

55 **Text S3: Year differences in ecological conditions**

56 To describe the differences in the ecological conditions across the two study years, we
57 considered average daily temperature and daily probability of rainfall, two variables previously
58 shown to influence the energetic requirements of great tit nestlings (Radford *et al.* 2001).
59 Weather data were obtained from the Rothenfeld weather station located in the center of the
60 study area (<http://www.gkd.bayern.de/>). We tested for between-year differences in average daily
61 temperature (°C) and total daily probability of precipitation (rain = 1, no rain = 0) throughout the
62 breeding season using a generalized linear model (glm). Data were included for the months of

63 April, May and June, to cover the period from egg laying to nestling fledging for our first broods.
64 In 2010, weather data was unavailable for 4 days during this period, resulting in N = 87 days of
65 weather data for 2010 and N = 91 days of weather data for 2011. Year, April day (days since
66 April 1st), and their interaction were included as fixed effects. Temperature was modelled with
67 Gaussian errors and probability of precipitation was modelled with binomial errors. Analyses
68 using daily precipitation (in mm) gave qualitatively similar results to analyses of daily
69 probability of rainfall, and we present only the latter analyses here.

70 We also investigated annual differences in nestling mortality, both before and after the brood
71 size manipulations. To control for annual differences in clutch size, we modelled mortality as a
72 two-column matrix comprised of the number of nestlings dead and the number of nestlings alive
73 (using the “cbind” function). For mortality prior to the manipulations, these values were the
74 number of nestlings that died between hatching and nestling day 3 (dead) and the number of
75 nestlings that survived until day 3 (alive). For mortality after the brood size manipulation, these
76 values were the number of nestlings that died between day 3 and fledging (dead), and the number
77 of nestlings that survived from day 3 until fledging (alive). For analyses of mortality prior to the
78 BSM, mortality was modelled as a function of year, and plot was included as a random effect.
79 For analyses of post-manipulation mortality, we modelled mortality as a function of year-specific
80 treatment (6-level factor: 2010 Reduced, 2010 Control, 2010 Enlarged, 2011 Reduced, 2011
81 Control, 2011 Enlarged). Models were fitted with a binomial error distribution (coded as family
82 = “multinomial2”).

83 Analyses of temperature and precipitation data confirmed earlier work suggesting that 2010 was
84 a more challenging season to great tits breeding in our study populations compared with 2011
85 (Nicolaus *et al.* 2015). At the start of the breeding season, average daily temperatures were 3.8°C

86 (95% CI = 1.68, 5.96) higher in 2011 compared with 2010. Average daily temperatures increased
87 across the breeding season (April to June) in both years (2010: $\beta = 0.13^{\circ}\text{C}$ per day, 95% CI =
88 0.09, 0.16; 2011: $\beta = 0.07^{\circ}\text{C}$, 95% CI = 0.04, 0.10, Fig. S2a), but the increase was greater in
89 2010. Thus, year differences in temperatures were greatest early in the breeding season (Fig.
90 S2a).

91 The 2010 breeding season was also more challenging with regards to precipitation (Fig. S2b). At
92 the beginning of the breeding season, there was moderate support for the interpretation that the
93 daily probability of precipitation was greater in 2010 compared with 2011 (odds ratio $\beta = 1.54$,
94 95% CI = -0.24, 2.90). The daily probability of precipitation increased in 2011 ($\beta = 0.04$, 95% CI
95 = 0.02, 0.05), but there was no evidence for a seasonal increase in 2010 ($\beta = 0.01$, 95% CI = -
96 0.01, 0.02). Thus, as with year differences in temperature, year differences in precipitation were
97 greatest early in the breeding season (Fig. S2b).

98 The year differences in ecological conditions were also reflected in year differences in nestling
99 mortality. Prior to manipulations, mortality was higher in 2010 compared with 2011 (mean %
100 mortality: 3.5% versus 0.2%, Fig. S3a). This corresponded to an estimated difference in odds
101 ratios of $\beta = 3.58$, 95% CI = 1.60, 5.70). Post-manipulation mortality was also higher in 2010
102 (Fig. S2b, mean % mortality: 55% in 2010 versus 31% in 2011). Additionally, mortality
103 increased across BSM categories in 2010 (i.e. enlarged broods had higher mortality than reduced
104 broods without overlap in 95% CIs), but not in 2011 (Table S1 and Fig. S3b).

105 **Table S1:** Variation in mean inter-visit interval (IVI, in seconds), load size (bill volumes), and prey type (preferred or non-preferred),
 106 and short-term delivery experienced at the level of nestlings (load size/IVI/brood size) in relation to brood size (within-year centered),
 107 year, and their interaction. IVI and delivery data were ln transformed, and load size was square-root transformed. Data were modelled
 108 with Gaussian (IVI, load size, delivery) or binomial (prey type) errors. Year-BSM-specific residual variances were fitted for IVI, load
 109 size and delivery, but not for prey type, as it is not possible to fit heterogeneous residual errors to binomial data. Values presented are
 110 estimated means (β) and their 95% credible intervals (CI).

Fixed effects	Ln IVI $\beta \pm 95\% \text{ CI}$	Sqrt load size $\beta \pm 95\% \text{ CI}$	Prey type¹ $\beta \pm 95\% \text{ CI}$	Ln delivery $\beta \pm 95\% \text{ CI}$
Intercept ²	4.85 (4.74, 4.95)	0.96 (0.93, 1.00)	-8.31 (-14.75, -1.76)	-6.12 (-6.23, -6.00)
Brood size	-0.05 (-0.09, 0.01)	0.01 (0.00, 0.03)	0.23 (-0.65, 1.39)	-0.03, -0.09, 0.02)
Year (2011)	0.12 (-0.01, 0.26)	0.19 (0.14, 0.24)	11.82 (2.25, 20.82)	-0.05 (-0.21, 0.09)
Brood size: Year (2011)	-0.05 (-0.11, 0.02)	-0.02 (-0.04, 0.00)	-1.12 (-2.61, 0.11)	0.02 (-0.04, 0.09)
Residual variance	$\sigma^2 \pm 95\% \text{ CI}$	$\sigma^2 \pm 95\% \text{ CI}$	$\sigma^2 \pm 95\% \text{ CI}$	$\sigma^2 \pm 95\% \text{ CI}$
2010				
Reduced	0.55 (0.46, 0.65)	0.054 (0.046, 0.062)	NA	0.88 (0.74, 1.03)
Control	0.56 (0.48, 0.64)	0.058 (0.049, 0.068)	NA	1.05 (0.88, 1.21)
Enlarged	0.60 (0.50, 0.69)	0.046 (0.040, 0.053)	NA	1.11 (0.96, 1.28)
2011				
Reduced	0.76 (0.60, 0.93)	0.09 (0.07, 0.10)	NA	1.05 (0.87, 1.26)
Control	0.63 (0.51, 0.76)	0.09 (0.08, 0.11)	NA	1.15 (0.96, 1.36)
Enlarged	0.57 (0.49, 0.63)	0.08 (0.07, 0.09)	NA	1.04 (0.90, 1.17)

112 1. Prey type was modelled with binomial errors. Preferred prey (i.e. caterpillars) were coded as 1, and non-preferred prey (i.e. others)
113 were coded as 0. Estimates are in ln (odds ratios): therefore, an estimate of 0 represents an equal proportion of preferred and non-
114 preferred prey in the diet; an estimate of 0.69 represents cases where preferred prey are twice as frequent as non-preferred prey, and an
115 estimate of -0.69 represents cases where non-preferred prey are provisioned twice as much as preferred prey.

116 2. Intercept estimated at within-year centered brood size of zero, and year = 2010.

Table S2: Nestling mortality as a function of year and BSM. Nestling mortality was modelled using as the ln odds ratio of dying versus surviving (dead, alive) between the brood size manipulation day and fledging. Values presented are estimated means (β) or variances (σ^2) and their 95% credible intervals (CI).

Mortality (odds ratio of dying)¹	
<i>Fixed effects</i>	$\beta \pm 95\% \text{ CI}$
2010	
Reduced	-0.48 (-1.65, 0.92)
Control	0.85 (-0.50, 2.27)
Enlarged	2.41 (1.10, 3.60)
2011	
Reduced	-1.93 (-3.27, -0.46)
Control	-1.81 (-3.26, 0.48)
Enlarged	-0.84 (-2.18, 0.40)
<i>Random effects</i>	$\sigma^2 \pm 95\% \text{ CI}$
Plot	1.55 (0.26, 3.43)
Residual	8.95 (5.63, 12.77)

¹Estimates are in ln (odds ratios): therefore, an estimate of 0 represents an equal proportion of nestlings that died and survived in the specified interval (i.e. 50% mortality); an estimate of 0.69 represents cases where dying is twice as likely as surviving, and an estimate of -0.69 represents cases where surviving is twice as likely as dying.

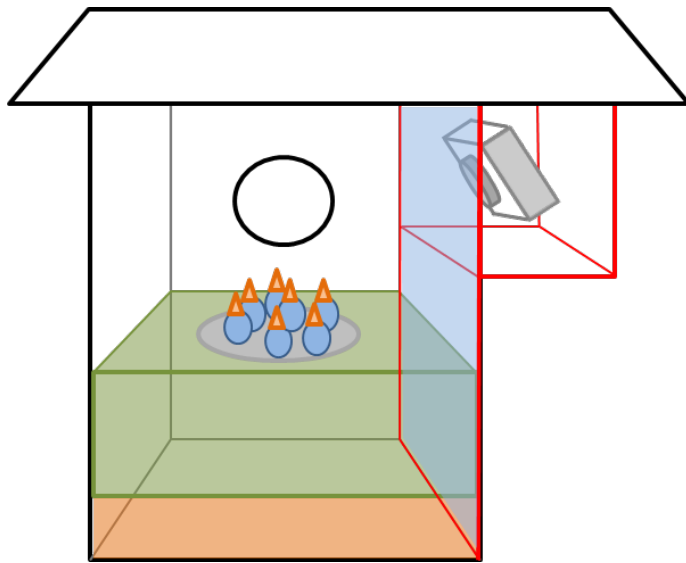


Figure S1: Schematic overview of nestbox during provisioning trials. The side door (illustrated in red) with camera (illustrated in grey) was installed two days before provisioning videos were recorded, and was replaced with the original door after recordings were complete. A sheet of clear plexiglass (illustrated in blue) was placed between the nest and the side door in order to prevent parents from sitting within the side compartment that housed the camera, which would have obscured views from the camera. Additionally, a small piece of foam (illustrated in orange) was placed below the nesting material (illustrated in green) to ensure that the full nestcup was visible in recordings.

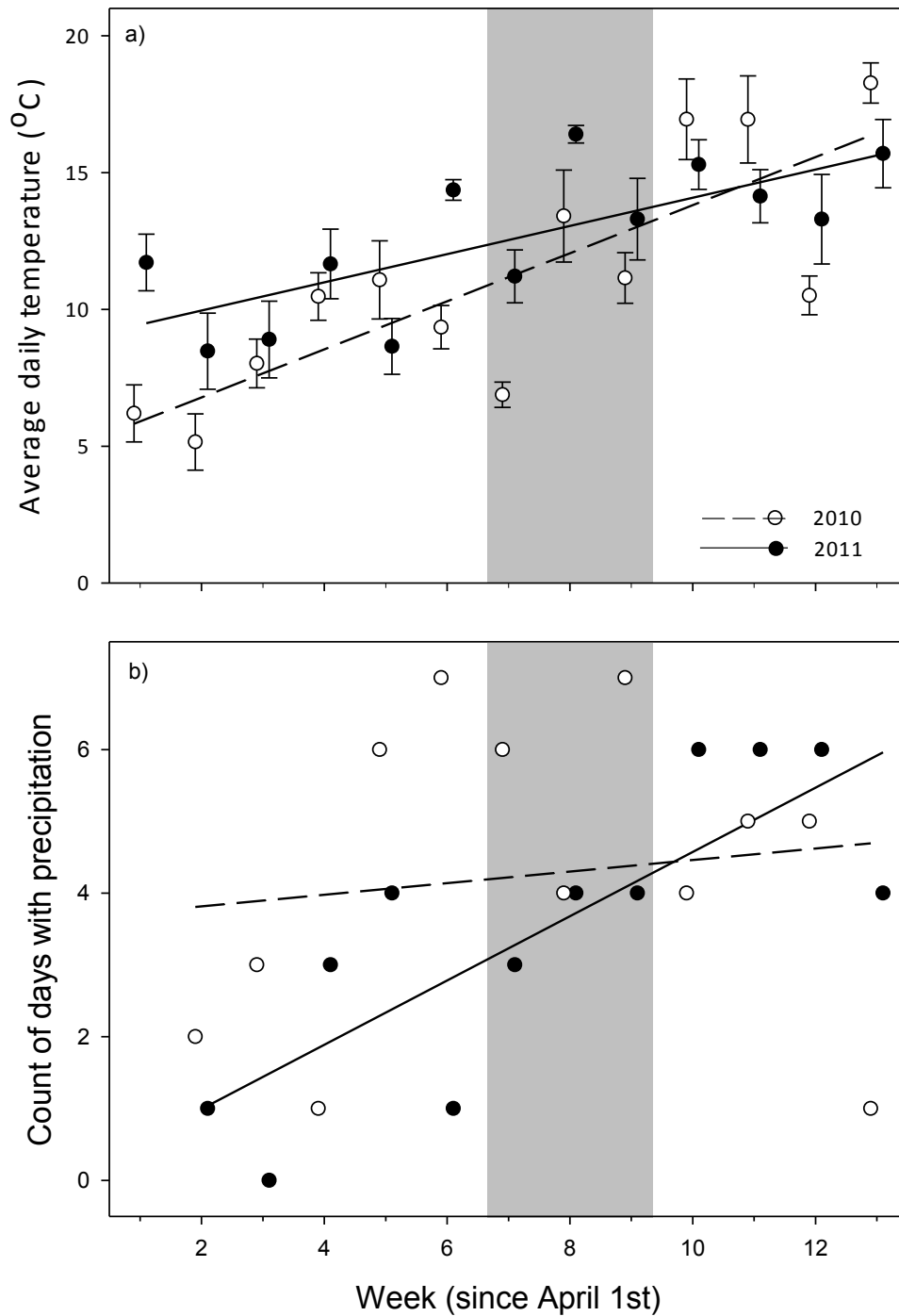


Figure S2: Annual differences in a) temperature and b) precipitation throughout the breeding seasons in 2010 and 2011. In a) values presented are weekly means \pm 1 s.e. In b) values presented are counts of days with precipitation in each week. Week 1 is not shown in Fig. 3b) because of missing data from the weather station. The weeks during which provisioning behavior was observed are indicated with grey shading. Lines are best fits to the raw data.

Supplementary References

- Nicolaus, M., Mathot, K.J., Araya-Ajoy, Y.G., Mutzel, A., Wijmenga, J.J., Kempenaers, B. & Dingemanse, N.J. (2015) Does coping style predict optimization? An experimental test in a wild passerine bird. *Proceedings of the Royal Society B Biological Sciences*, **282**, 20142405.
- Radford, A.N., McCleery, R.H., Woodburn, R.J.W. & Morecroft, M.D. (2001) Activity patterns of parent Great Tits *Parus major* feeding their young during rainfall. *Bird Study*, **48**, 214-220.